

# Large random matrices with given margins

---

Hanbaek Lyu

University of Wisconsin - Madison

Based on joint work with Sumit Mukherjee (Columbia)

Apr. 3, 2025

## Introduction

Random graphs with given degree sequences

A parametric approach for RMs with given margin

Contingency tables and Typical tables

A non-parametric approach to RMs with given margin

Some results on RMs with exactly given margins

Phase diagram of tame margins

Open problems

Sinkhorn algorithm

Static Shrödinger bridge

- ▶ (*Base model*)  $\mu$  = probability measure on  $\mathbb{R}$ , and let

$$A := \inf\{\text{supp}(\mu)\} \leq \sup\{\text{supp}(\mu)\} =: B.$$

$X \sim \mu^{\otimes(m \times n)}$ :  $(m \times n)$  random matrix with i.i.d. entries drawn from  $\mu$

- (*Base model*)  $\mu$  = probability measure on  $\mathbb{R}$ , and let

$$A := \inf\{\text{supp}(\mu)\} \leq \sup\{\text{supp}(\mu)\} =: B.$$

$X \sim \mu^{\otimes(m \times n)}$ :  $(m \times n)$  random matrix with i.i.d. entries drawn from  $\mu$

- (*Margins*) For a matrix  $\mathbf{x} = (x_{ij}) \in \mathbb{R}^{m \times n}$ ,  $(r(\mathbf{x}), c(\mathbf{x}))$  = margin of  $\mathbf{x}$ :

$$r(\mathbf{x}) := (r_1(\mathbf{x}), \dots, r_m(\mathbf{x})); \quad r_i(\mathbf{x}) := x_{i1} + \dots + x_{in} \quad (\triangleright \text{row margin of } \mathbf{x})$$

$$c(\mathbf{x}) := (c_1(\mathbf{x}), \dots, c_n(\mathbf{x})); \quad c_j(\mathbf{x}) := x_{1j} + \dots + x_{mj} \quad (\triangleright \text{column margin of } \mathbf{x})$$

$$\mathcal{T}_\rho(\mathbf{r}, \mathbf{c}) := \{\mathbf{x} \in \mathbb{R}^{m \times n} : \|(\mathbf{r}, \mathbf{c}) - (r(\mathbf{x}), c(\mathbf{x}))\|_1 \leq \rho\}$$

- (*Base model*)  $\mu$  = probability measure on  $\mathbb{R}$ , and let

$$A := \inf\{\text{supp}(\mu)\} \leq \sup\{\text{supp}(\mu)\} =: B.$$

$X \sim \mu^{\otimes(m \times n)}$ :  $(m \times n)$  random matrix with i.i.d. entries drawn from  $\mu$

- (*Margins*) For a matrix  $\mathbf{x} = (x_{ij}) \in \mathbb{R}^{m \times n}$ ,  $(r(\mathbf{x}), c(\mathbf{x}))$  = margin of  $\mathbf{x}$ :

$$r(\mathbf{x}) := (r_1(\mathbf{x}), \dots, r_m(\mathbf{x})); \quad r_i(\mathbf{x}) := x_{i1} + \dots + x_{in} \quad (\triangleright \text{row margin of } \mathbf{x})$$

$$c(\mathbf{x}) := (c_1(\mathbf{x}), \dots, c_n(\mathbf{x})); \quad c_j(\mathbf{x}) := x_{1j} + \dots + x_{mj} \quad (\triangleright \text{column margin of } \mathbf{x})$$

$$\mathcal{T}_\rho(\mathbf{r}, \mathbf{c}) := \{\mathbf{x} \in \mathbb{R}^{m \times n} : \|(\mathbf{r}, \mathbf{c}) - (r(\mathbf{x}), c(\mathbf{x}))\|_1 \leq \rho\}$$

- (*Main question*)

If we condition  $X \sim \mu^{\otimes(m \times n)}$  on being in  $\mathcal{T}_\rho(\mathbf{r}, \mathbf{c})$ , how does it look like?

► (Main question)

If we condition  $X \sim \mu^{\otimes (m \times n)}$  on being in  $\mathcal{T}_\rho(r, c)$ , how does it look like?

- $X_{11} \stackrel{d}{=} ?$
- $\mathbb{E}[X] = ?$
- $X - \mathbb{E}[X] \stackrel{d}{=} ?$
- $X \stackrel{d}{=} ?$

► (Main question)

If we condition  $X \sim \mu^{\otimes (m \times n)}$  on being in  $\mathcal{T}_\rho(r, c)$ , how does it look like?

- $X_{11} \stackrel{d}{=} ?$
- $\mathbb{E}[X] = ?$
- $X - \mathbb{E}[X] \stackrel{d}{=} ?$
- $X \stackrel{d}{=} ?$

► High-level answer:

► (Main question)

If we condition  $X \sim \mu^{\otimes (m \times n)}$  on being in  $\mathcal{T}_\rho(\mathbf{r}, \mathbf{c})$ , how does it look like?

- $X_{11} \stackrel{d}{=} ?$
- $\mathbb{E}[X] = ?$
- $X - \mathbb{E}[X] \stackrel{d}{=} ?$
- $X \stackrel{d}{=} ?$

► High-level answer:

- *Maximum Likelihood Perspective (Parameteric)*: The **maximum likelihood entrywise exponential tilting** of the base model for margin  $(\mathbf{r}, \mathbf{c})$



► (Main question)

If we condition  $X \sim \mu^{\otimes (m \times n)}$  on being in  $\mathcal{T}_\rho(\mathbf{r}, \mathbf{c})$ , how does it look like?

- $X_{11} \stackrel{d}{=} ?$
- $\mathbb{E}[X] = ?$
- $X - \mathbb{E}[X] \stackrel{d}{=} ?$
- $X \stackrel{d}{=} ?$

► High-level answer:

- *Maximum Likelihood Perspective (Parametric)*: The **maximum likelihood entrywise exponential tilting** of the base model for margin  $(\mathbf{r}, \mathbf{c})$
- *Minimum Relative Entropy Perspective (Non-parametric)*: The random matrix ensemble with **minimum relative entropy** from the base model constrained to have the expected margin  $(\mathbf{r}, \mathbf{c})$ .

► (Main question)

If we condition  $X \sim \mu^{\otimes (m \times n)}$  on being in  $\mathcal{T}_\rho(\mathbf{r}, \mathbf{c})$ , how does it look like?

- $X_{11} \stackrel{d}{=} ?$
- $\mathbb{E}[X] = ?$
- $X - \mathbb{E}[X] \stackrel{d}{=} ?$
- $X \stackrel{d}{=} ?$

► High-level answer:

- *Maximum Likelihood Perspective (Parameteric)*: The **maximum likelihood entrywise exponential tilting** of the base model for margin  $(\mathbf{r}, \mathbf{c})$
- *Minimum Relative Entropy Perspective (Non-parametric)*: The random matrix ensemble with **minimum relative entropy** from the base model constrained to have the expected margin  $(\mathbf{r}, \mathbf{c})$ .

► These two approaches give the same answer! (strong duality)

Introduction

Random graphs with given degree sequences

A parametric approach for RMs with given margin

Contingency tables and Typical tables

A non-parametric approach to RMs with given margin

Some results on RMs with exactly given margins

Phase diagram of tame margins

Open problems

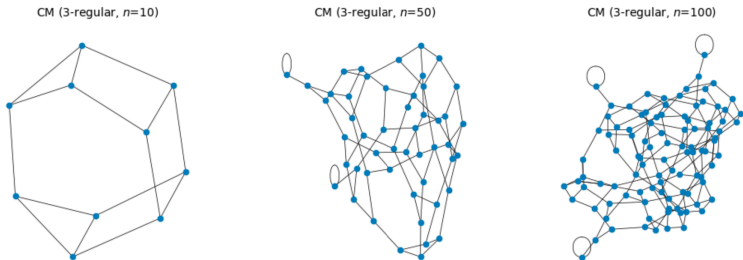
Sinkhorn algorithm

Static Shrödinger bridge

- ▶  $\mathbf{d} = (d_1, d_2, \dots, d_n)$ : **degree sequence** of an  $n$ -node graph

- ▶  $\mathbf{d} = (d_1, d_2, \dots, d_n)$ : **degree sequence** of an  $n$ -node graph
- ▶ (Question)

**How does a uniformly random graph with degree sequence  $\mathbf{d}$  look like?**



**Figure:** Random 3-regular graphs generated by the configuration model (allowing loops)

- ▶  $(\mathbf{d}^n)_{n \geq 1}$ : dense degree sequence with scaling limit to  $\mathbf{c} : [0, 1] \rightarrow (c_1, c_2) \subseteq (0, 1)$

- ▶  $(\mathbf{d}^n)_{n \geq 1}$ : dense degree sequence with scaling limit to  $\mathbf{c} : [0, 1] \rightarrow (c_1, c_2) \subseteq (0, 1)$
- ▶ Chatterjee, Diaconis, Sly '11 [5]

- ▶  $(\mathbf{d}^n)_{n \geq 1}$ : dense degree sequence with scaling limit to  $\mathbf{c} : [0, 1] \rightarrow (c_1, c_2) \subseteq (0, 1)$
- ▶ Chatterjee, Diaconis, Sly '11 [5]
  - Assume  $\mathbf{c}$  satisfies the 'continuum Erdős-Gallai condition'

$$\int_0^x \mathbf{c}(y) dy < x^2 + \int_x^1 \mathbf{c}(y) \wedge x dy$$



- ▶  $(\mathbf{d}^n)_{n \geq 1}$ : dense degree sequence with scaling limit to  $\mathbf{c} : [0, 1] \rightarrow (c_1, c_2) \subseteq (0, 1)$
- ▶ Chatterjee, Diaconis, Sly '11 [5]
  - Assume  $\mathbf{c}$  satisfies the 'continuum Erdős-Gallai condition'

$$\int_0^x \mathbf{c}(y) dy < x^2 + \int_x^1 \mathbf{c}(y) \wedge x dy$$

- There exists a limiting 'continuum dual variable'  $\beta^* : [0, 1] \rightarrow \mathbb{R}$  such that the corresponding graphon

$$W^{\beta^*}(x, y) = \frac{1}{\exp(\beta^*(x) + \beta^*(y)) + 1}$$

has 'degree sequence'  $\mathbf{c}$ :

$$\int_{\mathbb{R}} W^{\beta^*}(x, y) dy = \mathbf{c}(x)$$

- ▶  $(\mathbf{d}^n)_{n \geq 1}$ : dense degree sequence with scaling limit to  $\mathbf{c} : [0, 1] \rightarrow (c_1, c_2) \subseteq (0, 1)$
- ▶ Chatterjee, Diaconis, Sly '11 [5]
  - Assume  $\mathbf{c}$  satisfies the 'continuum Erdős-Gallai condition'

$$\int_0^x \mathbf{c}(y) dy < x^2 + \int_x^1 \mathbf{c}(y) \wedge x dy$$

- There exists a limiting 'continuum dual variable'  $\beta^* : [0, 1] \rightarrow \mathbb{R}$  such that the corresponding graphon

$$W^{\beta^*}(x, y) = \frac{1}{\exp(\beta^*(x) + \beta^*(y)) + 1}$$

has 'degree sequence'  $\mathbf{c}$ :

$$\int_{\mathbb{R}} W^{\beta^*}(x, y) dy = \mathbf{c}(x)$$

- $A^n = \text{Adj mx of the uniformly random graph with degree seq. } \mathbf{d}^n$ . Then

$$W_{A^n} \rightarrow W^{\beta^*} \quad \text{in weak cut distance a.s.,}$$

( $W_{A^n}$ : step function corresponding to the adj mx  $A^n$ )

- ▶ Fit a parametrized random graph model with independent edges (s.t. deg seq = sufficient statistic) to the target degree sequence by MLE

- ▶ Fit a **parametrized random graph model** with independent edges (s.t. deg seq = sufficient statistic) to the target degree sequence by MLE
  - **(The  $\beta$ -model)** Given a dual variable  $\beta \in \mathbb{R}^n$ ,  $G_\beta$  = random graph with  $n$  nodes and independent edges, where

$$\mathbb{P}_\beta(i \sim j) = \frac{e^{\beta(i)+\beta(j)}}{1 + e^{\beta(i)+\beta(j)}} = \mathbb{E}[A^\beta(i, j)]$$

- Fit a **parametrized random graph model** with independent edges (s.t. deg seq = sufficient statistic) to the target degree sequence by MLE
  - **(The  $\beta$ -model)** Given a dual variable  $\beta \in \mathbb{R}^n$ ,  $G_\beta$  = random graph with  $n$  nodes and independent edges, where

$$\mathbb{P}_\beta(i \sim j) = \frac{e^{\beta(i) + \beta(j)}}{1 + e^{\beta(i) + \beta(j)}} = \mathbb{E}[A^\beta(i, j)]$$

- **(log-likelihood)**

$$\ell(\beta) = \sum_{i,j} x_{ij}(\beta(i) + \beta(j)) - \underbrace{\log \left( 1 + e^{\beta(i) + \beta(j)} \right)}_{=\psi(\beta(i) + \beta(j))} = 2\langle \mathbf{d}, \beta \rangle - \sum_{i,j} \psi(\beta(i) + \beta(j))$$

- Fit a **parametrized random graph model** with independent edges (s.t. deg seq = sufficient statistic) to the target degree sequence by MLE
  - **(The  $\beta$ -model)** Given a dual variable  $\beta \in \mathbb{R}^n$ ,  $G_\beta$  = random graph with  $n$  nodes and independent edges, where

$$\mathbb{P}_\beta(i \sim j) = \frac{e^{\beta(i)+\beta(j)}}{1 + e^{\beta(i)+\beta(j)}} = \mathbb{E}[A^\beta(i, j)]$$

- **(log-likelihood)**

$$\ell(\beta) = \sum_{i,j} x_{ij}(\beta(i) + \beta(j)) - \underbrace{\log \left( 1 + e^{\beta(i)+\beta(j)} \right)}_{=\psi(\beta(i)+\beta(j))} = 2\langle \mathbf{d}, \beta \rangle - \sum_{i,j} \psi(\beta(i) + \beta(j))$$

- **(The MLE equation)**  $\frac{d\ell(\beta)}{d\beta} = 0 \iff \mathbb{E}[\text{degree seq.}] = \mathbf{d}$ :

$$\mathbb{E} \left[ \sum_{j=1}^n A^\beta(i, j) \right] = d_i \quad \text{for all } 1 \leq i \leq n$$

► **Sketch of proof:**

- Find MLE  $\beta^n$  for the  $\beta$ -model to the target degree sequence  $\mathbf{d}^n$

► **Sketch of proof:**

- Find MLE  $\beta^n$  for the  $\beta$ -model to the target degree sequence  $\mathbf{d}^n$
- Transference:

$$G^{\beta^n} \stackrel{d}{\approx} G(\mathbf{d}^n) \sim \text{Uniform}(G \text{ with deg seq} = \mathbf{d}^n)$$



► **Sketch of proof:**

- Find MLE  $\beta^n$  for the  $\beta$ -model to the target degree sequence  $\mathbf{d}^n$
- Transference:

$$G^{\beta^n} \stackrel{d}{\approx} G(\mathbf{d}^n) \sim \text{Uniform}(G \text{ with deg seq} = \mathbf{d}^n)$$

$$1. \ G^{\beta^n} \mid \{\text{deg seq} = \mathbf{d}^n\} \stackrel{d}{=} G(\mathbf{d}^n)$$

► **Sketch of proof:**

- Find MLE  $\beta^n$  for the  $\beta$ -model to the target degree sequence  $\mathbf{d}^n$
- Transference:

$$G^{\beta^n} \stackrel{d}{\approx} G(\mathbf{d}^n) \sim \text{Uniform}(G \text{ with deg seq} = \mathbf{d}^n)$$

1.  $G^{\beta^n} \mid \{\text{deg seq} = \mathbf{d}^n\} \stackrel{d}{=} G(\mathbf{d}^n)$
2.  $\mathbb{P}(G^{\beta^n} \text{ has deg seq } \mathbf{d}^n) \geq \exp(-o(n^{3/2+\varepsilon}))$

► **Sketch of proof:**

- Find MLE  $\beta^n$  for the  $\beta$ -model to the target degree sequence  $\mathbf{d}^n$
- Transference:

$$G^{\beta^n} \stackrel{d}{\approx} G(\mathbf{d}^n) \sim \text{Uniform}(G \text{ with deg seq } = \mathbf{d}^n)$$

1.  $G^{\beta^n} \mid \{\text{deg seq} = \mathbf{d}^n\} \stackrel{d}{=} G(\mathbf{d}^n)$
2.  $\mathbb{P}(G^{\beta^n} \text{ has deg seq } \mathbf{d}^n) \geq \exp(-o(n^{3/2+\varepsilon}))$

- The above implies

$$\begin{aligned} \mathbb{P}(G(\mathbf{d}^n) \in \mathcal{E}) &= \mathbb{P}\left(G^{\beta^n} \in \mathcal{E} \mid G^{\beta^n} \text{ has deg seq } \mathbf{d}^n\right) \\ &\leq \mathbb{P}(G^{\beta^n} \text{ has deg seq } \mathbf{d}^n)^{-1} \mathbb{P}\left(G^{\beta^n} \in \mathcal{E}\right) \\ &\leq \exp(o(n^{3/2+\varepsilon})) \mathbb{P}\left(G^{\beta^n} \in \mathcal{E}\right) \end{aligned}$$

► **Sketch of proof:**

- Find MLE  $\beta^n$  for the  $\beta$ -model to the target degree sequence  $\mathbf{d}^n$
- Transference:

$$G^{\beta^n} \stackrel{d}{\approx} G(\mathbf{d}^n) \sim \text{Uniform}(G \text{ with deg seq } = \mathbf{d}^n)$$

1.  $G^{\beta^n} \mid \{\text{deg seq} = \mathbf{d}^n\} \stackrel{d}{=} G(\mathbf{d}^n)$
2.  $\mathbb{P}(G^{\beta^n} \text{ has deg seq } \mathbf{d}^n) \geq \exp(-o(n^{3/2+\varepsilon}))$

- The above implies

$$\begin{aligned} \mathbb{P}(G(\mathbf{d}^n) \in \mathcal{E}) &= \mathbb{P}\left(G^{\beta^n} \in \mathcal{E} \mid G^{\beta^n} \text{ has deg seq } \mathbf{d}^n\right) \\ &\leq \mathbb{P}(G^{\beta^n} \text{ has deg seq } \mathbf{d}^n)^{-1} \mathbb{P}(G^{\beta^n} \in \mathcal{E}) \\ &\leq \exp(o(n^{3/2+\varepsilon})) \mathbb{P}(G^{\beta^n} \in \mathcal{E}) \end{aligned}$$

- So events extremely rare under  $G^{\beta^n}$  are also rare under  $G(\mathbf{d}^n)$

► **Sketch of proof:**

- Find MLE  $\beta^n$  for the  $\beta$ -model to the target degree sequence  $\mathbf{d}^n$
- Transference:

$$G^{\beta^n} \stackrel{d}{\approx} G(\mathbf{d}^n) \sim \text{Uniform}(G \text{ with deg seq } = \mathbf{d}^n)$$

- $G^{\beta^n} \mid \{\text{deg seq} = \mathbf{d}^n\} \stackrel{d}{=} G(\mathbf{d}^n)$
- $\mathbb{P}(G^{\beta^n} \text{ has deg seq } \mathbf{d}^n) \geq \exp(-o(n^{3/2+\varepsilon}))$

- The above implies

$$\begin{aligned} \mathbb{P}(G(\mathbf{d}^n) \in \mathcal{E}) &= \mathbb{P}\left(G^{\beta^n} \in \mathcal{E} \mid G^{\beta^n} \text{ has deg seq } \mathbf{d}^n\right) \\ &\leq \mathbb{P}(G^{\beta^n} \text{ has deg seq } \mathbf{d}^n)^{-1} \mathbb{P}(G^{\beta^n} \in \mathcal{E}) \\ &\leq \exp(o(n^{3/2+\varepsilon})) \mathbb{P}(G^{\beta^n} \in \mathcal{E}) \end{aligned}$$

- So events extremely rare under  $G^{\beta^n}$  are also rare under  $G(\mathbf{d}^n)$   
e.g.,  $G(\mathbf{d}^n)$  cannot be too far from  $\mathbb{E}[G^{\beta^n}] \approx G^{\beta^n}$

Introduction

Random graphs with given degree sequences

A parametric approach for RMs with given margin

Contingency tables and Typical tables

A non-parametric approach to RMs with given margin

Some results on RMs with exactly given margins

Phase diagram of tame margins

Open problems

Sinkhorn algorithm

Static Shrödinger bridge

- ▶ Goal: Approximate  $X \sim \mu^{\otimes(m \times n)}$  given  $X \in \mathcal{T}_\rho(\mathbf{r}, \mathbf{c})$

- ▶ Goal: Approximate  $X \sim \mu^{\otimes(m \times n)}$  given  $X \in \mathcal{T}_\rho(\mathbf{r}, \mathbf{c})$
- ▶  $\mu_\theta :=$  exponentially tilted probability measure given by

$$\frac{d\mu_\theta}{d\mu}(x) = e^{\theta x - \psi(\theta)}, \quad \psi(\theta) := \log \int_{\mathbb{R}} e^{\theta x} d\mu(x) = \log \text{partition function}$$



- ▶ Goal: Approximate  $X \sim \mu^{\otimes(m \times n)}$  given  $X \in \mathcal{T}_\rho(\mathbf{r}, \mathbf{c})$

- ▶  $\mu_\theta :=$  exponentially tilted probability measure given by

$$\frac{d\mu_\theta}{d\mu}(x) = e^{\theta x - \psi(\theta)}, \quad \psi(\theta) := \log \int_{\mathbb{R}} e^{\theta x} d\mu(x) = \log \text{partition function}$$

- ▶ Elementary facts:

$$\mathbb{E}_{X \sim \mu_\theta}[X] = \psi'(\theta), \quad \text{Var}_{X \sim \mu_\theta}(X) = \psi''(\theta) > 0.$$

- $\psi' : \{\text{tilting params.}\} \rightarrow (A, B)$  is strictly increasing ( $\triangleright$  tilt2mean function)
- $\phi = (\psi')^{-1} : (A, B) \rightarrow \{\text{tilting params.}\}$  is strictly increasing ( $\triangleright$  mean2tilt function)

- ▶ Goal: Approximate  $X \sim \mu^{\otimes(m \times n)}$  given  $X \in \mathcal{T}_\rho(\mathbf{r}, \mathbf{c})$

- ▶ Goal: Approximate  $X \sim \mu^{\otimes(m \times n)}$  given  $X \in \mathcal{T}_\rho(\mathbf{r}, \mathbf{c})$
- ▶ Given vectors  $\alpha, \beta$ ,  $Y = (Y_{ij}) \sim \mu_{\alpha \oplus \beta}$ : RM w/ indep. entries  $Y_{ij} \sim \mu_{\alpha(i) + \beta(j)}$

- ▶ Goal: Approximate  $X \sim \mu^{\otimes(m \times n)}$  given  $X \in \mathcal{T}_\rho(\mathbf{r}, \mathbf{c})$
- ▶ Given vectors  $\alpha, \beta$ ,  $Y = (Y_{ij}) \sim \mu_{\alpha \oplus \beta}$ : RM w/ indep. entries  $Y_{ij} \sim \mu_{\alpha(i) + \beta(j)}$
- ▶ log-likelihood of margin  $(\mathbf{r}, \mathbf{c})$  under  $Y$  w.r.t.  $\mu^{\otimes(m \times n)}$ :

$$\sum_{i,j} \left[ x_{ij}(\alpha(i) + \beta(j)) - \psi(\alpha(i) + \beta(j)) \right] = \langle \mathbf{r}, \alpha \rangle + \langle \mathbf{c}, \beta \rangle - \sum_{i,j} \psi(\alpha(i) + \beta(j)).$$

- ▶ Goal: Approximate  $X \sim \mu^{\otimes(m \times n)}$  given  $X \in \mathcal{T}_\rho(\mathbf{r}, \mathbf{c})$
- ▶ Given vectors  $\alpha, \beta$ ,  $Y = (Y_{ij}) \sim \mu_{\alpha \oplus \beta}$ : RM w/ indep. entries  $Y_{ij} \sim \mu_{\alpha(i) + \beta(j)}$
- ▶ log-likelihood of margin  $(\mathbf{r}, \mathbf{c})$  under  $Y$  w.r.t.  $\mu^{\otimes(m \times n)}$ :

$$\sum_{i,j} \left[ x_{ij}(\alpha(i) + \beta(j)) - \psi(\alpha(i) + \beta(j)) \right] = \langle \mathbf{r}, \alpha \rangle + \langle \mathbf{c}, \beta \rangle - \sum_{i,j} \psi(\alpha(i) + \beta(j)).$$

- ▶ MLE for margin  $(\mathbf{r}, \mathbf{c})$ :

$$\sup_{\alpha, \beta} \left( g^{\mathbf{r}, \mathbf{c}}(\alpha, \beta) := \langle \mathbf{r}, \alpha \rangle + \langle \mathbf{c}, \beta \rangle - \sum_{i,j} \psi(\alpha(i) + \beta(j)) \right),$$

- ▶ Goal: Approximate  $X \sim \mu^{\otimes(m \times n)}$  given  $X \in \mathcal{T}_\rho(\mathbf{r}, \mathbf{c})$
- ▶ Given vectors  $\alpha, \beta$ ,  $Y = (Y_{ij}) \sim \mu_{\alpha \oplus \beta}$ : RM w/ indep. entries  $Y_{ij} \sim \mu_{\alpha(i) + \beta(j)}$
- ▶ log-likelihood of margin  $(\mathbf{r}, \mathbf{c})$  under  $Y$  w.r.t.  $\mu^{\otimes(m \times n)}$ :

$$\sum_{i,j} \left[ x_{ij}(\alpha(i) + \beta(j)) - \psi(\alpha(i) + \beta(j)) \right] = \langle \mathbf{r}, \alpha \rangle + \langle \mathbf{c}, \beta \rangle - \sum_{i,j} \psi(\alpha(i) + \beta(j)).$$

- ▶ MLE for margin  $(\mathbf{r}, \mathbf{c})$ :

$$\sup_{\alpha, \beta} \left( g^{\mathbf{r}, \mathbf{c}}(\alpha, \beta) := \langle \mathbf{r}, \alpha \rangle + \langle \mathbf{c}, \beta \rangle - \sum_{i,j} \psi(\alpha(i) + \beta(j)) \right),$$

- ▶ Taking  $\nabla g^{\mathbf{r}, \mathbf{c}}(\alpha, \beta) = \mathbf{0}$ , get the MLE equation:

$$\mathbb{E}[Y] = \psi'(\alpha \oplus \beta) \in \mathcal{T}(\mathbf{r}, \mathbf{c})$$

- ▶ Goal: Approximate  $X \sim \mu^{\otimes(m \times n)}$  given  $X \in \mathcal{T}_\rho(\mathbf{r}, \mathbf{c})$
- ▶ Given vectors  $\alpha, \beta$ ,  $Y = (Y_{ij}) \sim \mu_{\alpha \oplus \beta}$ : RM w/ indep. entries  $Y_{ij} \sim \mu_{\alpha(i) + \beta(j)}$
- ▶ log-likelihood of margin  $(\mathbf{r}, \mathbf{c})$  under  $Y$  w.r.t.  $\mu^{\otimes(m \times n)}$ :

$$\sum_{i,j} \left[ x_{ij}(\alpha(i) + \beta(j)) - \psi(\alpha(i) + \beta(j)) \right] = \langle \mathbf{r}, \alpha \rangle + \langle \mathbf{c}, \beta \rangle - \sum_{i,j} \psi(\alpha(i) + \beta(j)).$$

- ▶ MLE for margin  $(\mathbf{r}, \mathbf{c})$ :

$$\sup_{\alpha, \beta} \left( g^{\mathbf{r}, \mathbf{c}}(\alpha, \beta) := \langle \mathbf{r}, \alpha \rangle + \langle \mathbf{c}, \beta \rangle - \sum_{i,j} \psi(\alpha(i) + \beta(j)) \right),$$

- ▶ Taking  $\nabla g^{\mathbf{r}, \mathbf{c}}(\alpha, \beta) = \mathbf{0}$ , get the MLE equation:

$$\mathbb{E}[Y] = \psi'(\alpha \oplus \beta) \in \mathcal{T}(\mathbf{r}, \mathbf{c})$$

- MLE is not unique:  $(\alpha, \beta) \iff (\alpha + \lambda, \beta - \lambda)$

- ▶ Goal: Approximate  $X \sim \mu^{\otimes(m \times n)}$  given  $X \in \mathcal{T}_\rho(\mathbf{r}, \mathbf{c})$
- ▶ Given vectors  $\alpha, \beta$ ,  $Y = (Y_{ij}) \sim \mu_{\alpha \oplus \beta}$ : RM w/ indep. entries  $Y_{ij} \sim \mu_{\alpha(i) + \beta(j)}$
- ▶ log-likelihood of margin  $(\mathbf{r}, \mathbf{c})$  under  $Y$  w.r.t.  $\mu^{\otimes(m \times n)}$ :

$$\sum_{i,j} \left[ x_{ij}(\alpha(i) + \beta(j)) - \psi(\alpha(i) + \beta(j)) \right] = \langle \mathbf{r}, \alpha \rangle + \langle \mathbf{c}, \beta \rangle - \sum_{i,j} \psi(\alpha(i) + \beta(j)).$$

- ▶ MLE for margin  $(\mathbf{r}, \mathbf{c})$ :

$$\sup_{\alpha, \beta} \left( g^{\mathbf{r}, \mathbf{c}}(\alpha, \beta) := \langle \mathbf{r}, \alpha \rangle + \langle \mathbf{c}, \beta \rangle - \sum_{i,j} \psi(\alpha(i) + \beta(j)) \right),$$

- ▶ Taking  $\nabla g^{\mathbf{r}, \mathbf{c}}(\alpha, \beta) = \mathbf{0}$ , get the MLE equation:

$$\mathbb{E}[Y] = \psi'(\alpha \oplus \beta) \in \mathcal{T}(\mathbf{r}, \mathbf{c})$$

- MLE is not unique:  $(\alpha, \beta) \iff (\alpha + \lambda, \beta - \lambda)$
- How do we compute an MLE? No closed form; use Sinkhorn-type algorithm (will revisit)



- (*Informal result I*: Maximum likelihood perspective)

$$\left[ X \sim \mu^{\otimes(m \times n)} \text{ conditioned on being in } \mathcal{T}_\rho(r, c) \right] \approx Y \sim \mu_{\alpha \oplus \beta},$$

where  $(\alpha, \beta)$  is an MLE for  $(r, c)$

- ▶ Behavior of an  $(\alpha, \beta)$ -model depends crucially on how far the entries of  $\alpha \oplus \beta$  are away from the extreme values  $\phi(A)$  and  $\phi(B)$

- Behavior of an  $(\alpha, \beta)$ -model depends crucially on how far the entries of  $\alpha \oplus \beta$  are away from the extreme values  $\phi(A)$  and  $\phi(B)$

### Definition (Tame margins)

An  $m \times n$  margin  $(\mathbf{r}, \mathbf{c})$  is  $\delta$ -tame for  $\delta > 0$  if the MLE  $(\alpha, \beta)$  exists and its entries satisfy (recall  $(A, B) = \text{Int}(\text{supp}(\mu))$ )

$$A_\delta := \max \left( A + \delta, -\frac{1}{\delta} \right) \leq \psi'(\alpha \oplus \beta) \leq \min \left( B - \delta, \frac{1}{\delta} \right) =: B_\delta.$$

- Behavior of an  $(\alpha, \beta)$ -model depends crucially on how far the entries of  $\alpha \oplus \beta$  are away from the extreme values  $\phi(A)$  and  $\phi(B)$

### Definition (Tame margins)

An  $m \times n$  margin  $(\mathbf{r}, \mathbf{c})$  is  $\delta$ -tame for  $\delta > 0$  if the MLE  $(\alpha, \beta)$  exists and its entries satisfy (recall  $(A, B) = \text{Int}(\text{supp}(\mu))$ )

$$A_\delta := \max\left(A + \delta, -\frac{1}{\delta}\right) \leq \psi'(\alpha \oplus \beta) \leq \min\left(B - \delta, \frac{1}{\delta}\right) =: B_\delta.$$

- If an MLE  $(\alpha, \beta)$  exists for  $(\mathbf{r}, \mathbf{c})$ , then  $(\mathbf{r}, \mathbf{c})$  is always  $\delta$ -tame for some  $\delta > 0$  that may depend on  $m$  and  $n$ .

- ▶ Behavior of an  $(\alpha, \beta)$ -model depends crucially on how far the entries of  $\alpha \oplus \beta$  are away from the extreme values  $\phi(A)$  and  $\phi(B)$

### Definition (Tame margins)

An  $m \times n$  margin  $(\mathbf{r}, \mathbf{c})$  is  $\delta$ -tame for  $\delta > 0$  if the MLE  $(\alpha, \beta)$  exists and its entries satisfy (recall  $(A, B) = \text{Int}(\text{supp}(\mu))$ )

$$A_\delta := \max \left( A + \delta, -\frac{1}{\delta} \right) \leq \psi'(\alpha \oplus \beta) \leq \min \left( B - \delta, \frac{1}{\delta} \right) =: B_\delta.$$

- ▶ If an MLE  $(\alpha, \beta)$  exists for  $(\mathbf{r}, \mathbf{c})$ , then  $(\mathbf{r}, \mathbf{c})$  is always  $\delta$ -tame for some  $\delta > 0$  that may depend on  $m$  and  $n$ .
- ▶ A technical issue: When is a sequence of margins **uniformly**  $\delta$ -tame? (will revisit)

## Theorem (Transference; L-M '24+)

$(\mathbf{r}, \mathbf{c}) = m \times n$   $\delta$ -tame margin with an MLE  $(\alpha, \beta)$ , and let  $X \sim \mu^{\otimes(m \times n)}$  be conditional on  $X \in \mathcal{T}_\rho(\mathbf{r}, \mathbf{c})$  for some  $\rho \geq 0$ . Let  $Y \sim \mu_{\alpha \oplus \beta}$ . Then for some constant  $C = C(\mu, \delta) > 0$ , for each measurable set  $\mathcal{E} \subseteq \mathbb{R}^{m \times n}$ ,

## Theorem (Transference; L-M '24+)

$(\mathbf{r}, \mathbf{c}) = m \times n$   $\delta$ -tame margin with an MLE  $(\alpha, \beta)$ , and let  $X \sim \mu^{\otimes(m \times n)}$  be conditional on  $X \in \mathcal{T}_\rho(\mathbf{r}, \mathbf{c})$  for some  $\rho \geq 0$ . Let  $Y \sim \mu_{\alpha \oplus \beta}$ . Then for some constant  $C = C(\mu, \delta) > 0$ , for each measurable set  $\mathcal{E} \subseteq \mathbb{R}^{m \times n}$ ,

$$\mathbb{P}(X \in \mathcal{E}) \leq \exp(C\rho) \mathbb{P}(Y \in \mathcal{T}_\rho(\mathbf{r}, \mathbf{c}))^{-1} \mathbb{P}(Y \in \mathcal{E}).$$

## Theorem (Transference; L-M '24+)

$(\mathbf{r}, \mathbf{c}) = m \times n$   $\delta$ -tame margin with an MLE  $(\alpha, \beta)$ , and let  $X \sim \mu^{\otimes(m \times n)}$  be conditional on  $X \in \mathcal{T}_\rho(\mathbf{r}, \mathbf{c})$  for some  $\rho \geq 0$ . Let  $Y \sim \mu_{\alpha \oplus \beta}$ . Then for some constant  $C = C(\mu, \delta) > 0$ , for each measurable set  $\mathcal{E} \subseteq \mathbb{R}^{m \times n}$ ,

$$\mathbb{P}(X \in \mathcal{E}) \leq \exp(C\rho) \mathbb{P}(Y \in \mathcal{T}_\rho(\mathbf{r}, \mathbf{c}))^{-1} \mathbb{P}(Y \in \mathcal{E}).$$

In particular, if  $\rho \geq \sqrt{mn(m+n)}$ , then for each  $t \geq 0$ ,

$$\mathbb{P}(X \in \mathcal{E}) \leq 2 \exp(C\rho) \mathbb{P}(Y \in \mathcal{E}).$$



## Theorem (Transference; L-M '24+)

$(\mathbf{r}, \mathbf{c}) = m \times n$   $\delta$ -tame margin with an MLE  $(\alpha, \beta)$ , and let  $X \sim \mu^{\otimes(m \times n)}$  be conditional on  $X \in \mathcal{T}_\rho(\mathbf{r}, \mathbf{c})$  for some  $\rho \geq 0$ . Let  $Y \sim \mu_{\alpha \oplus \beta}$ . Then for some constant  $C = C(\mu, \delta) > 0$ , for each measurable set  $\mathcal{E} \subseteq \mathbb{R}^{m \times n}$ ,

$$\mathbb{P}(X \in \mathcal{E}) \leq \exp(C\rho) \mathbb{P}(Y \in \mathcal{T}_\rho(\mathbf{r}, \mathbf{c}))^{-1} \mathbb{P}(Y \in \mathcal{E}).$$

In particular, if  $\rho \geq \sqrt{mn(m+n)}$ , then for each  $t \geq 0$ ,

$$\mathbb{P}(X \in \mathcal{E}) \leq 2 \exp(C\rho) \mathbb{P}(Y \in \mathcal{E}).$$

- Events extremely rare under  $Y$  are also rare under  $X$

## Theorem (Transference; L-M '24+)

$(\mathbf{r}, \mathbf{c}) = m \times n$   $\delta$ -tame margin with an MLE  $(\alpha, \beta)$ , and let  $X \sim \mu^{\otimes(m \times n)}$  be conditional on  $X \in \mathcal{T}_\rho(\mathbf{r}, \mathbf{c})$  for some  $\rho \geq 0$ . Let  $Y \sim \mu_{\alpha \oplus \beta}$ . Then for some constant  $C = C(\mu, \delta) > 0$ , for each measurable set  $\mathcal{E} \subseteq \mathbb{R}^{m \times n}$ ,

$$\mathbb{P}(X \in \mathcal{E}) \leq \exp(C\rho) \mathbb{P}(Y \in \mathcal{T}_\rho(\mathbf{r}, \mathbf{c}))^{-1} \mathbb{P}(Y \in \mathcal{E}).$$

In particular, if  $\rho \geq \sqrt{mn(m+n)}$ , then for each  $t \geq 0$ ,

$$\mathbb{P}(X \in \mathcal{E}) \leq 2 \exp(C\rho) \mathbb{P}(Y \in \mathcal{E}).$$

- Events extremely rare under  $Y$  are also rare under  $X$
- $\mathbb{P}(Y \in \mathcal{T}_\rho(\mathbf{r}, \mathbf{c})) \geq 1/2$  for  $\rho \sim \sqrt{mn(m+n)}$  ( $\mathbb{E}[Y] \in \mathcal{T}(\mathbf{r}, \mathbf{c})$  and use concentration for  $Y$ )

- ▶ A *kernel* is an integrable function  $W: [0, 1]^2 \rightarrow \mathbb{R}$ . The *cut-norm* of a kernel  $W$  is defined as

$$\|W\|_{\square} := \sup_{S, T \subseteq [0, 1]} \left| \int_{S \times T} W(x, y) dx dy \right|.$$

Given an  $m \times n$  matrix  $A$ , define a step-kernel  $W_A$  as

$$W_A(x, y) := A_{ij} \text{ if } (x, y) \in R_{ij} = \left( \frac{i-1}{m}, \frac{i}{m} \right] \times \left( \frac{j-1}{n}, \frac{j}{n} \right]$$

- ▶ A *kernel* is an integrable function  $W: [0, 1]^2 \rightarrow \mathbb{R}$ . The *cut-norm* of a kernel  $W$  is defined as

$$\|W\|_{\square} := \sup_{S, T \subseteq [0, 1]} \left| \int_{S \times T} W(x, y) dx dy \right|.$$

Given an  $m \times n$  matrix  $A$ , define a step-kernel  $W_A$  as

$$W_A(x, y) := A_{ij} \text{ if } (x, y) \in R_{ij} = \left( \frac{i-1}{m}, \frac{i}{m} \right] \times \left( \frac{j-1}{n}, \frac{j}{n} \right]$$

## Theorem (Concentration in cut norm; L-M '24+ )

Keep the same setting as before. Then there exists a constant  $C = C(\delta, \mu) > 0$  s.t.

$$\mathbb{P}(\|W_X - W_{\mathbb{E}[Y]}\|_{\square} \geq t) \leq \underbrace{\mathbb{P}(Y \in \mathcal{T}_{\rho}(\mathbf{r}, \mathbf{c}))^{-1}}_{\text{transference cost}} \underbrace{\exp\left(C\rho + (m+n+1)\log 2 - \frac{t^2 mn}{C}\right)}_{\text{Concentration of } \|Y - \mathbb{E}[Y]\|_{\square}}.$$

Introduction

Random graphs with given degree sequences

A parametric approach for RMs with given margin

Contingency tables and Typical tables

A non-parametric approach to RMs with given margin

Some results on RMs with exactly given margins

Phase diagram of tame margins

Open problems

Sinkhorn algorithm

Static Shrödinger bridge

- **Contingency tables** = matrices with non-negative integer entries with fixed row and column margins

<i>Data</i>							
1	0	3	2	0	7		13
1	2	0	4	3	0		10
7	5	2	1	0	0		15
0	0	3	1	3	9		16
0	3	1	8	0	2		14
5	3	0	3	5	3		19
9	13	9	19	11	21		

*v. s.*

<i>Null model</i>							
$X = (X_{ij})$							13
							10
							15
							16
							14
							19
9	13	9	19	11	21		

- Contingency tables are fundamental tools in statistics for studying dependence structure between two or more variables
- Uniform contingency table  $X = (X_{ij})$  serves as the maximum entropy null model given margins

## Conjecture (Independence heuristic, Good '50)

$$|\mathcal{T}(\mathbf{r}, \mathbf{c})| \approx G(\mathbf{r}, \mathbf{c})$$

where

$$G(\mathbf{r}, \mathbf{c}) := \binom{N + mn - 1}{mn - 1}^{-1} \prod_{i=1}^m \binom{\mathbf{r}(i) + n - 1}{n - 1} \prod_{j=1}^n \binom{\mathbf{c}(j) + m - 1}{m - 1}.$$

Conjecture (Independence heuristic, Good '50)

$$|\mathcal{T}(\mathbf{r}, \mathbf{c})| \approx G(\mathbf{r}, \mathbf{c})$$

where

$$G(\mathbf{r}, \mathbf{c}) := \binom{N + mn - 1}{mn - 1}^{-1} \prod_{i=1}^m \binom{\mathbf{r}(i) + n - 1}{n - 1} \prod_{j=1}^n \binom{\mathbf{c}(j) + m - 1}{m - 1}.$$

**Good says:** “A random table with total sum  $N$  independently satisfies the row and column margins”



## Conjecture (Independence heuristic, Good '50)

$$|\mathcal{T}(\mathbf{r}, \mathbf{c})| \approx G(\mathbf{r}, \mathbf{c})$$

where

$$G(\mathbf{r}, \mathbf{c}) := \binom{N + mn - 1}{mn - 1}^{-1} \prod_{i=1}^m \binom{\mathbf{r}(i) + n - 1}{n - 1} \prod_{j=1}^n \binom{\mathbf{c}(j) + m - 1}{m - 1}.$$

**Good says:** “A random table with total sum  $N$  independently satisfies the row and column margins”

- $X \sim \text{Uniform}(\mathcal{S}_N)$ ,  $\mathcal{S}_N := \{\text{CT's with total sum } N = \sum \mathbf{r}(i) = \sum \mathbf{c}(j)\}$
- $\mathcal{R}_n(\mathbf{r}) := \{X \text{ has row margins } \mathbf{r}\}$ ,  $\mathcal{C}_m(\mathbf{c}) := \{X \text{ has column margins } \mathbf{c}\}$ .
- $\mathbb{P}(\mathcal{R}_n(\mathbf{r}) \cap \mathcal{C}_m(\mathbf{c})) = \frac{|\mathcal{T}(\mathbf{r}, \mathbf{c})|}{|\mathcal{S}_N|}$ ,  $\mathbb{P}(\mathcal{R}_n(\mathbf{r})) = \frac{|\mathcal{R}_n(\mathbf{r})|}{|\mathcal{S}_N|}$ ,  $\mathbb{P}(\mathcal{C}_m(\mathbf{c})) = \frac{|\mathcal{C}_m(\mathbf{c})|}{|\mathcal{S}_N|}$
- $|\mathcal{S}_N| = \binom{N + mn - 1}{mn - 1}$ ,  $|\mathcal{R}_n(\mathbf{r})| = \prod_{i=1}^m \binom{\mathbf{r}(i) + n - 1}{n - 1}$ ,  $|\mathcal{C}_m(\mathbf{c})| = \prod_{j=1}^n \binom{\mathbf{c}(j) + m - 1}{m - 1}$
- $$\frac{\mathbb{P}(\mathcal{R}_n(\mathbf{r}) \cap \mathcal{C}_m(\mathbf{c}))}{\mathbb{P}(\mathcal{R}_n(\mathbf{r})) \mathbb{P}(\mathcal{C}_m(\mathbf{c}))} = \frac{|\mathcal{T}(\mathbf{r}, \mathbf{c})|}{G(\mathbf{r}, \mathbf{c})}$$

## History of the Independence Heuristic (IH) $|\mathcal{T}(\mathbf{r}, \mathbf{c})| \approx G(\mathbf{a}, \mathbf{b})$ :

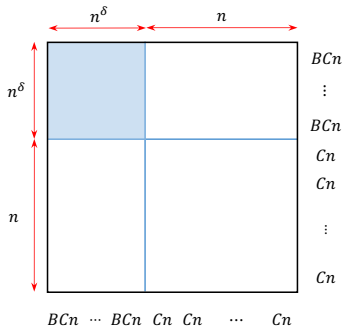
- Given implicitly by Good in 1963 [13] and later formally in 1963 [11] and 1976 [12]
- Experimentally verified by Good and Crook [10] in 1977 and Diagonis and Gangolli [7] in 1995
- Canfield and McKay '10 [4]: For  $m = n$  and  $\mathbf{r} = \mathbf{c} = (\lfloor Cn \rfloor, \dots, \lfloor Cn \rfloor)$ ,

$$\begin{aligned} \log |\mathcal{T}(\mathbf{r}, \mathbf{c})| &= [(1 + C) \log(1 + C) - C \log(C)]n^2 - n \log n \\ &\quad - n \log 2\pi C(1 + C) + \log n + O(1) \\ &\sim \log \sqrt{e} G(\mathbf{r}, \mathbf{c}) \end{aligned}$$

- In 2008, Greenhill and McKay [14] proved same asymptotics for **uniform but sparse margins**:  $\max(\mathbf{r}) \cdot \max(\mathbf{c}) = O(N^{2/3})$

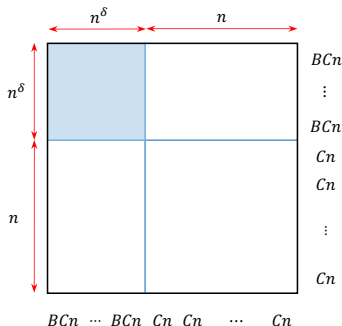
## But what about non-uniform margins?

- $2 \times 2$  block (Barvinok) margins:  $\mathbf{r} = \mathbf{c} = (\overbrace{BCn, \dots, BCn}^{n^\delta}, \overbrace{Cn, \dots, Cn}^{(n-n^\delta)}), 0 \leq \delta \leq 1$



## But what about non-uniform margins?

- $2 \times 2$  block (Barvinok) margins:  $\mathbf{r} = \mathbf{c} = (\overbrace{BCn, \dots, BCn}^{n^\delta}, \overbrace{Cn, \dots, Cn}^{(n-n^\delta)}), 0 \leq \delta \leq 1$



- **IH undercounts:** For  $\delta = 1$ , Barvinok [1] shows that

$$\lim_{n \rightarrow \infty} \frac{1}{n^2} \log |\mathcal{T}(\mathbf{r}, \mathbf{c})| > \lim_{n \rightarrow \infty} \frac{1}{n^2} \log G(\mathbf{r}, \mathbf{c}).$$

In other words, the rows and columns of CTs **attract** each other

- ▶ Barvinok '10: “To count  $|\mathcal{T}(\mathbf{r}, \mathbf{c})|$ , need to better understand  $\text{Uniform}(\mathcal{T}(\mathbf{r}, \mathbf{c}))$ ”

- ▶ Barvinok '10: “To count  $|\mathcal{T}(\mathbf{r}, \mathbf{c})|$ , need to better understand  $\text{Uniform}(\mathcal{T}(\mathbf{r}, \mathbf{c}))$ ”
- ▶ (Barvinok '10 [2]) For a  $m \times n$  margin  $(\mathbf{r}, \mathbf{c})$ , the corresponding typical table is

$$Z^{\mathbf{r}, \mathbf{c}} := \arg \max_{Q \in \mathcal{T}(\mathbf{r}, \mathbf{c})} \left[ g(Z) = \sum_{i,j} \underbrace{(z_{ij} + 1) \log(z_{ij} + 1) - z_{ij} \log(z_{ij})}_{= \text{Entropy}(\text{Geom}(\text{mean} = z_{ij}))} \right]$$

- ▶ Barvinok '10: “To count  $|\mathcal{T}(\mathbf{r}, \mathbf{c})|$ , need to better understand  $\text{Uniform}(\mathcal{T}(\mathbf{r}, \mathbf{c}))$ ”
- ▶ (Barvinok '10 [2]) For a  $m \times n$  margin  $(\mathbf{r}, \mathbf{c})$ , the corresponding typical table is

$$Z^{\mathbf{r}, \mathbf{c}} := \arg \max_{Q \in \mathcal{T}(\mathbf{r}, \mathbf{c})} \left[ g(Z) = \sum_{i,j} \underbrace{(z_{ij} + 1) \log(z_{ij} + 1) - z_{ij} \log(z_{ij})}_{=\text{Entropy}(\text{Geom}(\text{mean} = z_{ij}))} \right]$$

- ▶ Barvinok's insight:

$$\text{Uniform}(\mathcal{T}(\mathbf{r}, \mathbf{c})) \approx Z^{\mathbf{r}, \mathbf{c}}$$

- ▶ Barvinok '10: “To count  $|\mathcal{T}(\mathbf{r}, \mathbf{c})|$ , need to better understand  $\text{Uniform}(\mathcal{T}(\mathbf{r}, \mathbf{c}))$ ”
- ▶ (Barvinok '10 [2]) For a  $m \times n$  margin  $(\mathbf{r}, \mathbf{c})$ , the corresponding **typical table** is

$$Z^{\mathbf{r}, \mathbf{c}} := \arg \max_{Q \in \mathcal{T}(\mathbf{r}, \mathbf{c})} \left[ g(Z) = \sum_{i,j} \underbrace{(z_{ij} + 1) \log(z_{ij} + 1) - z_{ij} \log(z_{ij})}_{= \text{Entropy}(\text{Geom}(\text{mean} = z_{ij}))} \right]$$

- ▶ Barvinok's insight:

$$\text{Uniform}(\mathcal{T}(\mathbf{r}, \mathbf{c})) \approx Z^{\mathbf{r}, \mathbf{c}}$$

- (Barvinok '09 [1], '10 [2])

$$g(Z^{\mathbf{r}, \mathbf{c}}) - \gamma n \log n \leq \log |\mathcal{T}(\mathbf{r}, \mathbf{c})| \leq g(Z^{\mathbf{r}, \mathbf{c}})$$



- ▶ Barvinok '10: “To count  $|\mathcal{T}(\mathbf{r}, \mathbf{c})|$ , need to better understand  $\text{Uniform}(\mathcal{T}(\mathbf{r}, \mathbf{c}))$ ”
- ▶ (Barvinok '10 [2]) For a  $m \times n$  margin  $(\mathbf{r}, \mathbf{c})$ , the corresponding **typical table** is

$$Z^{\mathbf{r}, \mathbf{c}} := \arg \max_{Q \in \mathcal{T}(\mathbf{r}, \mathbf{c})} \left[ g(Z) = \sum_{i,j} \underbrace{(z_{ij} + 1) \log(z_{ij} + 1) - z_{ij} \log(z_{ij})}_{=\text{Entropy}(\text{Geom}(\text{mean} = z_{ij}))} \right]$$

- ▶ Barvinok's insight:

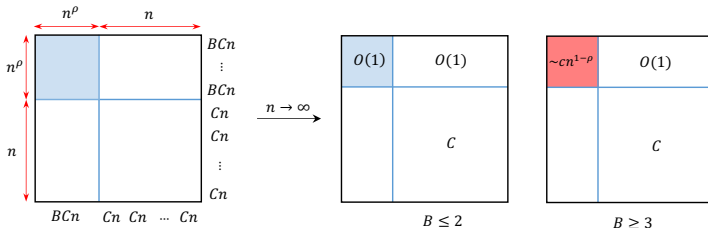
$$\text{Uniform}(\mathcal{T}(\mathbf{r}, \mathbf{c})) \approx Z^{\mathbf{r}, \mathbf{c}}$$

- (Barvinok '09 [1], '10 [2])

$$g(Z^{\mathbf{r}, \mathbf{c}}) - \gamma n \log n \leq \log |\mathcal{T}(\mathbf{r}, \mathbf{c})| \leq g(Z^{\mathbf{r}, \mathbf{c}})$$

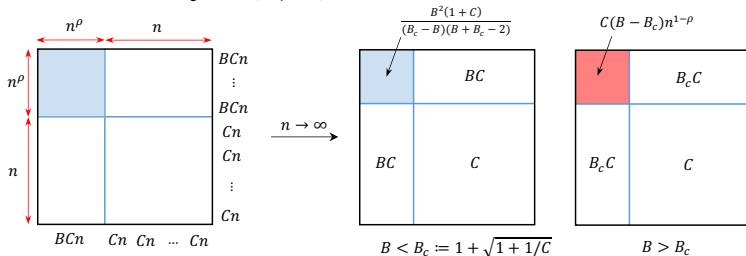
- Brändén, Leake, and Pak '23 [3] generalized this result to CTs with possibly bounded integer values (Using Lorenzian polynomials)

- In 2010, Barvinok conjectured that there is a phase transition in  $\text{Uniform}(\mathcal{T}(\text{Barv. margin}))$  as  $B$  increases



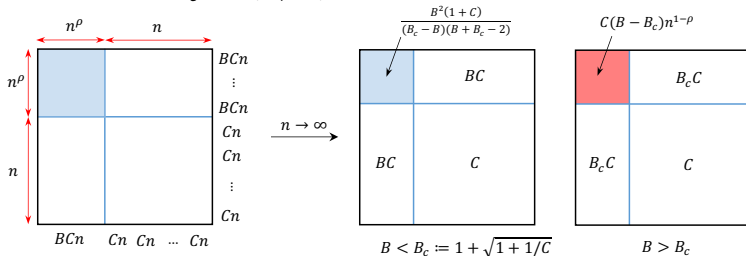
► Typical tables can change drastically by a small change in the margin!

- For  $0 \leq \delta < 1$ , Dittmer, Lyu, and Pak [8] show that  $Z^{r,c}$  undergoes a **sharp phase transition** at  $B_c = 1 + \sqrt{1 + C^{-1}}$ :



► Typical tables can change drastically by a small change in the margin!

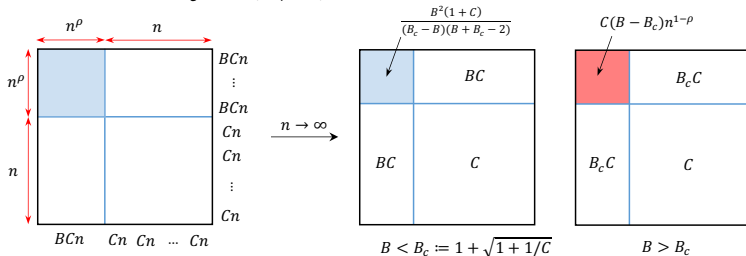
- For  $0 \leq \delta < 1$ , Dittmer, Lyu, and Pak [8] show that  $Z^{r,c}$  undergoes a **sharp phase transition** at  $B_c = 1 + \sqrt{1 + C^{-1}}$ :



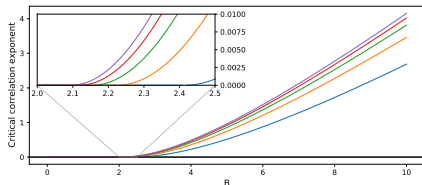
- $\delta$ -tame for  $B < B_c$ , non-tame for  $B > B_c$

► Typical tables can change drastically by a small change in the margin!

- For  $0 \leq \delta < 1$ , Dittmer, Lyu, and Pak [8] show that  $Z^{r,c}$  undergoes a **sharp phase transition** at  $B_c = 1 + \sqrt{1 + C^{-1}}$ :



- $\delta$ -tame for  $B < B_c$ , non-tame for  $B > B_c$
- This result was used to obtain a second-order phase transition in the number of CTs with Barvinok margin by Lyu and Pak '22 [15] ( $\log |\mathcal{T}(\mathbf{r}, \mathbf{c})| \approx g(Z^{r,c})$ )



Asymptotic independence  $\xrightarrow{B \nearrow}$   
Positive correlation

Introduction

Random graphs with given degree sequences

A parametric approach for RMs with given margin

Contingency tables and Typical tables

**A non-parametric approach to RMs with given margin**

Some results on RMs with exactly given margins

Phase diagram of tame margins

Open problems

Sinkhorn algorithm

Static Shrödinger bridge

- ▶ Goal: Approximate  $X \sim \mu^{\otimes(m \times n)}$  given  $X \in \mathcal{T}_\rho(\mathbf{r}, \mathbf{c})$

- ▶ Goal: Approximate  $X \sim \mu^{\otimes(m \times n)}$  given  $X \in \mathcal{T}_\rho(\mathbf{r}, \mathbf{c})$
- ▶ For each probability measure  $\mathcal{H}$  on  $\mathbb{R}^{m \times n}$ , the *relative entropy of  $\mathcal{H}$  from  $\mathcal{R}$*  is

$$D_{KL}(\mathcal{H} \parallel \mathcal{R}) := \int_{\mathbf{x} \in \mathbb{R}^{m \times n}} \frac{d\mathcal{H}(\mathbf{x})}{d\mathcal{R}(\mathbf{x})} \log \left( \frac{d\mathcal{H}(\mathbf{x})}{d\mathcal{R}(\mathbf{x})} \right) \mathcal{R}(d\mathbf{x}) \quad \text{if } \mathcal{H} \ll \mathcal{R} \text{ and } \infty \text{ o/w,}$$



- ▶ Goal: Approximate  $X \sim \mu^{\otimes(m \times n)}$  given  $X \in \mathcal{T}_\rho(\mathbf{r}, \mathbf{c})$
- ▶ For each probability measure  $\mathcal{H}$  on  $\mathbb{R}^{m \times n}$ , the *relative entropy of  $\mathcal{H}$  from  $\mathcal{R}$*  is

$$D_{KL}(\mathcal{H} \parallel \mathcal{R}) := \int_{\mathbf{x} \in \mathbb{R}^{m \times n}} \frac{d\mathcal{H}(\mathbf{x})}{d\mathcal{R}(\mathbf{x})} \log \left( \frac{d\mathcal{H}(\mathbf{x})}{d\mathcal{R}(\mathbf{x})} \right) \mathcal{R}(d\mathbf{x}) \quad \text{if } \mathcal{H} \ll \mathcal{R} \text{ and } \infty \text{ o/w,}$$

- ▶ Minimum relative entropy principle (information projection)

$$\left[ X \sim \mu^{\otimes(m \times n)} \text{ given } X \in \mathcal{T}_\rho(\mathbf{r}, \mathbf{c}) \right]$$
$$\stackrel{d}{\approx} \arg \min_{\mathcal{H} \in \mathcal{P}^{m \times n}} D_{KL}(\mathcal{H} \parallel \mathcal{R}) \quad \text{subject to} \quad \mathbb{E}_{X \sim \mathcal{H}}[(r(X), c(X))] = (\mathbf{r}, \mathbf{c})$$

- ▶ Goal: Approximate  $X \sim \mu^{\otimes(m \times n)}$  given  $X \in \mathcal{T}_\rho(\mathbf{r}, \mathbf{c})$
- ▶ For each probability measure  $\mathcal{H}$  on  $\mathbb{R}^{m \times n}$ , the *relative entropy of  $\mathcal{H}$  from  $\mathcal{R}$*  is

$$D_{KL}(\mathcal{H} \parallel \mathcal{R}) := \int_{\mathbf{x} \in \mathbb{R}^{m \times n}} \frac{d\mathcal{H}(\mathbf{x})}{d\mathcal{R}(\mathbf{x})} \log \left( \frac{d\mathcal{H}(\mathbf{x})}{d\mathcal{R}(\mathbf{x})} \right) \mathcal{R}(d\mathbf{x}) \quad \text{if } \mathcal{H} \ll \mathcal{R} \text{ and } \infty \text{ o/w,}$$

- ▶ Minimum relative entropy principle (information projection)

$$\left[ X \sim \mu^{\otimes(m \times n)} \text{ given } X \in \mathcal{T}_\rho(\mathbf{r}, \mathbf{c}) \right]$$

$$\stackrel{d}{\approx} \arg \min_{\mathcal{H} \in \mathcal{P}^{m \times n}} D_{KL}(\mathcal{H} \parallel \mathcal{R}) \quad \text{subject to} \quad \mathbb{E}_{X \sim \mathcal{H}}[(r(X), c(X))] = (\mathbf{r}, \mathbf{c})$$

$$= \bigotimes_{i,j} \mu_{\phi(z_{ij})} \quad \text{where} \quad Z = \arg \min_{Q=(q_{ij}) \in \mathcal{T}(\mathbf{r}, \mathbf{c})} \underbrace{D_{KL} \left( \bigotimes_{i,j} \mu_{\phi(q_{ij})} \parallel \mu^{\otimes(m \times n)} \right)}_{= \sum_{i,j} D_{KL}(\mu_{\phi(q_{ij})} \parallel \mu)}$$

- ▶ The **relative entropy** from the base measure  $\mu$  to the tilted probability measure  $\mu_\theta$ :

$$D(\mu_\theta \| \mu) := \int_{x \in \mathbb{R}} \log \left( \frac{d\mu_\theta}{d\mu}(x) \right) d\mu_\theta(x) = \theta \psi'(\theta) - \psi(\theta).$$

- ▶ The **relative entropy** from the base measure  $\mu$  to the tilted probability measure  $\mu_\theta$ :

$$D(\mu_\theta \| \mu) := \int_{x \in \mathbb{R}} \log \left( \frac{d\mu_\theta}{d\mu}(x) \right) d\mu_\theta(x) = \theta \psi'(\theta) - \psi(\theta).$$

- ▶ Fix a  $m \times n$  margin  $(\mathbf{r}, \mathbf{c}) \in \mathbb{R}^m \times \mathbb{R}^n$ . The **typical table**  $Z$  for margin  $(\mathbf{r}, \mathbf{c})$  is

$$Z^{\mathbf{r}, \mathbf{c}} := \arg \min_{X \in \mathcal{T}(\mathbf{r}, \mathbf{c})} \sum_{i,j} \underbrace{D(\mu_{\phi(x_{ij})} \| \mu)}_{f(x) := D(\mu_{\phi(x)} \| \mu) = x \phi(x) - \psi(\phi(x))}$$

- Strictly convex objective since  $f(x) = \phi(x)$ ,  $f'(x) = \phi'(x) = \frac{1}{\text{Var}(\mu_{\phi(x)})} > 0$
- So the typical table  $Z^{\mathbf{r}, \mathbf{c}}$  is unique if it exists

- ▶ The **relative entropy** from the base measure  $\mu$  to the tilted probability measure  $\mu_\theta$ :

$$D(\mu_\theta \| \mu) := \int_{x \in \mathbb{R}} \log \left( \frac{d\mu_\theta}{d\mu}(x) \right) d\mu_\theta(x) = \theta \psi'(\theta) - \psi(\theta).$$

- ▶ Fix a  $m \times n$  margin  $(\mathbf{r}, \mathbf{c}) \in \mathbb{R}^m \times \mathbb{R}^n$ . The **typical table**  $Z$  for margin  $(\mathbf{r}, \mathbf{c})$  is

$$Z^{\mathbf{r}, \mathbf{c}} := \arg \min_{X \in \mathcal{T}(\mathbf{r}, \mathbf{c})} \sum_{i,j} \underbrace{D(\mu_{\phi(x_{ij})} \| \mu)}_{f(x) := D(\mu_{\phi(x)} \| \mu) = x \phi(x) - \psi(\phi(x))}$$

- Strictly convex objective since  $f(x) = \phi(x)$ ,  $f'(x) = \phi'(x) = \frac{1}{\text{Var}(\mu_{\phi(x)})} > 0$
- So the typical table  $Z^{\mathbf{r}, \mathbf{c}}$  is unique if it exists
- ▶ By multivariate Lagrange multipliers, there are 'dual variables'  $\alpha \in \mathbb{R}^m$ ,  $\beta \in \mathbb{R}^n$  s.t.

$$Z^{\mathbf{r}, \mathbf{c}} = \psi'(\alpha \oplus \beta) = \mathbb{E}[\mu_{\alpha \oplus \beta}]!!$$

- Dual variable  $(\alpha, \beta)$  determined by the margin condition: (MLE!!)

$$\sum_{i=1}^m \psi'(\alpha(i) + \beta(j)) = \mathbf{r}(i), \quad \sum_{j=1}^n \psi'(\alpha(i) + \beta(j)) = \mathbf{c}(j) \quad \forall i, j$$

*(Informal result II: Minimum relative entropy perspective)*

$\left[ X \sim \mu^{\otimes(m \times n)} \text{ conditioned on being in } \mathcal{T}(\mathbf{r}, \mathbf{c}) \right] \approx \text{typical table } Z^{\mathbf{r}, \mathbf{c}}$   
where  $Z^{\mathbf{r}, \mathbf{c}} = \psi'(\alpha \oplus \beta)$  for some  $\alpha \in \mathbb{R}^m, \beta \in \mathbb{R}^n$

- $\mu = \text{Gaussian}$

$$\Theta = \mathbb{R}, \quad (A, B) = (-\infty, \infty), \quad \psi(\theta) = \frac{\theta^2}{2}, \quad \psi'(\theta) = \theta, \quad \phi(x) = x$$

$$f(x) = x\phi(x) - \psi(\phi(x)) = \frac{x^2}{2}$$

$$Z_{ij}^{\mathbf{r}, \mathbf{c}} = \frac{\mathbf{r}(i)}{n} + \frac{\mathbf{c}(j)}{m} - \frac{N}{mn} \quad (N = \sum_i \mathbf{r}(i) = \sum_j \mathbf{c}(j))$$

►  $\mu = \text{Gaussian}$

$$\Theta = \mathbb{R}, \quad (A, B) = (-\infty, \infty), \quad \psi(\theta) = \frac{\theta^2}{2}, \quad \psi'(\theta) = \theta, \quad \phi(x) = x$$

$$f(x) = x\phi(x) - \psi(\phi(x)) = \frac{x^2}{2}$$

$$Z_{ij}^{\mathbf{r}, \mathbf{c}} = \frac{\mathbf{r}(i)}{n} + \frac{\mathbf{c}(j)}{m} - \frac{N}{mn} \quad (N = \sum_i \mathbf{r}(i) = \sum_j \mathbf{c}(j))$$

►  $\mu = \text{Poisson}$

$$\Theta = \mathbb{R}, \quad (A, B) = (0, \infty), \quad \psi(\theta) = e^\theta, \quad \psi'(\theta) = e^\theta, \quad \phi(x) = \log x$$

$$f(x) = x\phi(x) - \psi(\phi(x)) = x \log x - x$$

$$Z_{ij}^{\mathbf{r}, \mathbf{c}} = e^{\alpha(i) + \beta(j)} = \mathbf{r}(i)\mathbf{c}(j)/N \quad (\triangleright \text{Fisher-Yates table})$$



►  $\mu = \text{Bernoulli}(1/2)$

$$\Theta = \mathbb{R}, \quad (A, B) = (0, 1), \quad \psi(\theta) = \log \frac{1 + e^\theta}{2}, \quad \psi'(\theta) = \frac{e^\theta}{1 + e^\theta}, \quad \phi(x) = \log \frac{x}{1 - x}$$

$$f(x) = x\phi(x) - \psi(\phi(x)) = x \log x + (1 - x) \log(1 - x) \quad \triangleright -\text{Entropy}(\text{Ber}(x))$$

$$Z_{ij}^{\mathbf{r}, \mathbf{c}} = \frac{1}{\exp(-\alpha(i) - \beta(j)) + 1} \quad \text{s.t. } Z^{\mathbf{r}, \mathbf{c}} \in \mathcal{T}(\mathbf{r}, \mathbf{c})$$

- $\mu = \text{Bernoulli}(1/2)$

$$\Theta = \mathbb{R}, \quad (A, B) = (0, 1), \quad \psi(\theta) = \log \frac{1 + e^\theta}{2}, \quad \psi'(\theta) = \frac{e^\theta}{1 + e^\theta}, \quad \phi(x) = \log \frac{x}{1 - x}$$

$$f(x) = x\phi(x) - \psi(\phi(x)) = x \log x + (1 - x) \log(1 - x) \quad \triangleright -\text{Entropy}(\text{Ber}(x))$$

$$Z_{ij}^{\mathbf{r}, \mathbf{c}} = \frac{1}{\exp(-\alpha(i) - \beta(j)) + 1} \quad \text{s.t. } Z^{\mathbf{r}, \mathbf{c}} \in \mathcal{T}(\mathbf{r}, \mathbf{c})$$

- $\mu = \text{Counting}(\mathbb{Z}_{\geq 0})$

$$\Theta = (-\infty, 0), \quad \psi(\theta) = -\log(1 - e^\theta), \quad \psi'(\theta) = \frac{e^\theta}{1 - e^\theta}, \quad \phi(x) = -\log(1 + x^{-1})$$

$$f(x) = x\phi(x) - \psi(\phi(x)) = x \log x - (1 + x) \log(1 + x) \quad \triangleright -\text{Entropy}(\text{Geom}(x))$$

$$Z_{ij}^{\mathbf{r}, \mathbf{c}} = \frac{1}{\exp(-\alpha(i) - \beta(j)) - 1} \quad \text{s.t. } Z^{\mathbf{r}, \mathbf{c}} \in \mathcal{T}(\mathbf{r}, \mathbf{c})$$

## Theorem (Strong duality; L-M '24+)

$$\begin{aligned}
\arg \min_Z \textit{Relative Entropy}(Z) &= \arg \max_{\alpha, \beta} \textit{Likelihood Of Margin}(\alpha, \beta) \\
\arg \min_{Z \in \mathcal{T}(\mathbf{r}, \mathbf{c})} \sum_{i,j} D(\mu_{\phi(z_{ij})} \parallel \mu) &= \arg \max_{\alpha, \beta} \left( \langle \mathbf{r}, \alpha \rangle + \langle \mathbf{c}, \beta \rangle - \sum_{i,j} \psi(\alpha(i) + \beta(j)) \right) \\
\underbrace{Z_{ij}^{\mathbf{r}, \mathbf{c}}}_{\text{typical table}} &= \psi' \left( \underbrace{\alpha(i) + \beta(j)}_{=\text{MLE}} \right) \quad (\psi' = \text{tilt2mean function})
\end{aligned}$$

Static Schrödinger Bridge b/w  $\mathbf{r}, \mathbf{c} = \textit{Kantorovich dual with potential}(\alpha, \beta)$

Non-parametric = Parametric

- ▶ We know  $X \approx Z^{\mathbf{r}_m, \mathbf{c}_n}$  in  $\|\cdot\|_{\square}$ . Does the typical tables (and the MLEs) have scaling limit as  $(\mathbf{r}_m, \mathbf{c}_n) \rightarrow (\mathbf{r}, \mathbf{c})$ ?

- ▶ We know  $X \approx Z^{\mathbf{r}_m, \mathbf{c}_n}$  in  $\|\cdot\|_{\square}$ . Does the typical tables (and the MLEs) have scaling limit as  $(\mathbf{r}_m, \mathbf{c}_n) \rightarrow (\mathbf{r}, \mathbf{c})$ ?
- ▶ A **continuum margin**  $(\mathbf{r}, \mathbf{c}) =$  integrable functions  $\mathbf{r}, \mathbf{c} : (0, 1] \rightarrow \mathbb{R}$  such that  $\int_0^1 \mathbf{r}(x) dx = \int_0^1 \mathbf{c}(y) dy$

- ▶ We know  $X \approx Z^{\mathbf{r}_m, \mathbf{c}_n}$  in  $\|\cdot\|_{\square}$ . Does the typical tables (and the MLEs) have scaling limit as  $(\mathbf{r}_m, \mathbf{c}_n) \rightarrow (\mathbf{r}, \mathbf{c})$ ?
- ▶ A **continuum margin**  $(\mathbf{r}, \mathbf{c}) =$  integrable functions  $\mathbf{r}, \mathbf{c} : (0, 1] \rightarrow \mathbb{R}$  such that  $\int_0^1 \mathbf{r}(x) dx = \int_0^1 \mathbf{c}(y) dy$
- ▶ For a  $m \times n$  discrete margin  $(\mathbf{r}_m, \mathbf{c}_n)$ , define the corresponding **continuum step margin**  $(\bar{\mathbf{r}}_m, \bar{\mathbf{c}}_n)$  as

$$\bar{\mathbf{r}}_m(t) := n^{-1} \mathbf{r}_m(\lceil mt \rceil), \quad \bar{\mathbf{c}}_n(t) := m^{-1} \mathbf{c}_n(\lceil nt \rceil).$$

- ▶ We know  $X \approx Z^{\mathbf{r}_m, \mathbf{c}_n}$  in  $\|\cdot\|_{\square}$ . Does the typical tables (and the MLEs) have scaling limit as  $(\mathbf{r}_m, \mathbf{c}_n) \rightarrow (\mathbf{r}, \mathbf{c})$ ?
- ▶ A **continuum margin**  $(\mathbf{r}, \mathbf{c}) =$  integrable functions  $\mathbf{r}, \mathbf{c} : (0, 1] \rightarrow \mathbb{R}$  such that  $\int_0^1 \mathbf{r}(x) dx = \int_0^1 \mathbf{c}(y) dy$
- ▶ For a  $m \times n$  discrete margin  $(\mathbf{r}_m, \mathbf{c}_n)$ , define the corresponding **continuum step margin**  $(\bar{\mathbf{r}}_m, \bar{\mathbf{c}}_n)$  as

$$\bar{\mathbf{r}}_m(t) := n^{-1} \mathbf{r}_m(\lceil mt \rceil), \quad \bar{\mathbf{c}}_n(t) := m^{-1} \mathbf{c}_n(\lceil nt \rceil).$$

- ▶ A seq. of  $m \times n$  margins  $(\mathbf{r}_m, \mathbf{c}_n)$  **converges in  $L^1$**  to a continuum margin  $(\mathbf{r}, \mathbf{c})$  if

$$\lim_{m, n \rightarrow \infty} \|\mathbf{r} - \bar{\mathbf{r}}_m\|_1 + \|\mathbf{c} - \bar{\mathbf{c}}_n\|_1 = 0.$$

- ▶ We know  $X \approx Z^{\mathbf{r}_m, \mathbf{c}_n}$  in  $\|\cdot\|_{\square}$ . Does the typical tables (and the MLEs) have scaling limit as  $(\mathbf{r}_m, \mathbf{c}_n) \rightarrow (\mathbf{r}, \mathbf{c})$ ?
- ▶ A **continuum margin**  $(\mathbf{r}, \mathbf{c}) =$  integrable functions  $\mathbf{r}, \mathbf{c} : (0, 1] \rightarrow \mathbb{R}$  such that  $\int_0^1 \mathbf{r}(x) dx = \int_0^1 \mathbf{c}(y) dy$
- ▶ For a  $m \times n$  discrete margin  $(\mathbf{r}_m, \mathbf{c}_n)$ , define the corresponding **continuum step margin**  $(\bar{\mathbf{r}}_m, \bar{\mathbf{c}}_n)$  as

$$\bar{\mathbf{r}}_m(t) := n^{-1} \mathbf{r}_m(\lceil mt \rceil), \quad \bar{\mathbf{c}}_n(t) := m^{-1} \mathbf{c}_n(\lceil nt \rceil).$$

- ▶ A seq. of  $m \times n$  margins  $(\mathbf{r}_m, \mathbf{c}_n)$  **converges in  $L^1$**  to a continuum margin  $(\mathbf{r}, \mathbf{c})$  if

$$\lim_{m, n \rightarrow \infty} \|\mathbf{r} - \bar{\mathbf{r}}_m\|_1 + \|\mathbf{c} - \bar{\mathbf{c}}_n\|_1 = 0.$$

- ▶ (*Informal result III*)

**For  $(\mathbf{r}_m, \mathbf{c}_n) \rightarrow (\mathbf{r}, \mathbf{c})$  in  $L^1$  and  $X \sim \mu^{\otimes(m \times n)}$  conditioned on  $\mathcal{T}(\mathbf{r}_m, \mathbf{c}_n)$ ,**

**$W_X \rightarrow W^{\mathbf{r}, \mathbf{c}}$  a.s. in cut norm**

**where  $W^{\mathbf{r}, \mathbf{c}}(x, y) = \psi'(\alpha(x) + \beta(y))$  for some  $\alpha, \beta \in [0, 1] \rightarrow \mathbb{R}$ .**



## Theorem (Scaling limit of typical tables and MLEs; L-M '24+)

Fix  $\delta > 0$  and let  $(\mathbf{r}_m, \mathbf{c}_n)$  be a sequence of  $m \times n$   $\delta$ -tame margins converging to a continuum margin  $(\mathbf{r}, \mathbf{c})$  in  $L^1$  as  $m, n \rightarrow \infty$ . Then  $\exists$  bounded measurable functions  $\alpha, \beta : [0, 1] \rightarrow \mathbb{R}$  s.t.  $\int \alpha(x) dx = 0$  and the kernel

$$W^{\mathbf{r}, \mathbf{c}}(x, y) := \psi'(\alpha(x) + \beta(y))$$

has continuum margin  $(\mathbf{r}, \mathbf{c})$ .

Furthermore,

$$\begin{aligned} \|W^{\mathbf{r}, \mathbf{c}} - W_{\mathbf{Z}^m, \mathbf{c}_n}\|_2^2 &\leq C_\delta \|(\mathbf{r}, \mathbf{c}) - (\bar{\mathbf{r}}_m, \bar{\mathbf{c}}_n)\|_1 \\ \|\alpha - \bar{\alpha}_m\|_2^2 + \|\beta - \bar{\beta}_n\|_2^2 &\leq C_\delta \|(\mathbf{r}, \mathbf{c}) - (\bar{\mathbf{r}}_m, \bar{\mathbf{c}}_n)\|_1. \end{aligned}$$

Introduction

Random graphs with given degree sequences

A parametric approach for RMs with given margin

Contingency tables and Typical tables

A non-parametric approach to RMs with given margin

**Some results on RMs with exactly given margins**

Phase diagram of tame margins

Open problems

Sinkhorn algorithm

Static Shrödinger bridge

- ▶ Recall the transference principle for exact margin conditioning ( $\rho = 0$ ):

$$\mathbb{P}(X \in \mathcal{E}) \leq \mathbb{P}(Y \in \mathcal{T}(\mathbf{r}, \mathbf{c}))^{-1} \mathbb{P}(Y \in \mathcal{E}), \quad Y \sim \mu_{\alpha \oplus \beta}$$

- ▶ Recall the transference principle for exact margin conditioning ( $\rho = 0$ ):

$$\mathbb{P}(X \in \mathcal{E}) \leq \mathbb{P}(Y \in \mathcal{T}(\mathbf{r}, \mathbf{c}))^{-1} \mathbb{P}(Y \in \mathcal{E}), \quad Y \sim \mu_{\alpha \oplus \beta}$$

- ▶ The key issue is to lower bound  $\mathbb{P}(Y \in \mathcal{T}(\mathbf{r}, \mathbf{c}))$

- ▶ Recall the transference principle for exact margin conditioning ( $\rho = 0$ ):

$$\mathbb{P}(X \in \mathcal{E}) \leq \mathbb{P}(Y \in \mathcal{T}(\mathbf{r}, \mathbf{c}))^{-1} \mathbb{P}(Y \in \mathcal{E}), \quad Y \sim \mu_{\alpha \oplus \beta}$$

- ▶ The key issue is to lower bound  $\mathbb{P}(Y \in \mathcal{T}(\mathbf{r}, \mathbf{c}))$

## Corollary (L-M '24+)

Let  $X \sim \mu^{\otimes(m \times n)}$  cond. on  $X \in \mathcal{T}(\mathbf{r}, \mathbf{c})$ . Suppose  $(\mathbf{r}, \mathbf{c})$  is a 'k-cloning' of some  $m_0 \times n_0$  margin  $(\mathbf{a}, \mathbf{b})$  with an MLE  $(\alpha_0, \beta_0)$ , i.e.,  $\mathbf{r} = \mathbf{a} \otimes \mathbf{1}_k$  and  $\mathbf{c} = \mathbf{b} \otimes \mathbf{1}_k$ , where  $\otimes$  denotes the Kronecker product. Then as  $k \rightarrow \infty$ ,

$$d_{TV}(X_{11}, \mu_{\alpha_0(1) + \beta_0(1)}) = \begin{cases} O\left(k^{-1/2} \sqrt{\log k}\right) & \text{if } \mu = \text{Counting}(\mathbb{Z}_{\geq 0}) \text{ or } \text{Leb}(\mathbb{R}_{\geq 0}) \\ O\left(k^{-1/4} \log k\right) & \text{if } \mu = \text{Wtd versions of the above} \end{cases}$$

- ▶ Recall the transference principle for exact margin conditioning ( $\rho = 0$ ):

$$\mathbb{P}(X \in \mathcal{E}) \leq \mathbb{P}(Y \in \mathcal{T}(\mathbf{r}, \mathbf{c}))^{-1} \mathbb{P}(Y \in \mathcal{E}), \quad Y \sim \mu_{\alpha \oplus \beta}$$

- ▶ The key issue is to lower bound  $\mathbb{P}(Y \in \mathcal{T}(\mathbf{r}, \mathbf{c}))$

## Corollary (L-M '24+)

Let  $X \sim \mu^{\otimes(m \times n)}$  cond. on  $X \in \mathcal{T}(\mathbf{r}, \mathbf{c})$ . Suppose  $(\mathbf{r}, \mathbf{c})$  is a ' $k$ -cloning' of some  $m_0 \times n_0$  margin  $(\mathbf{a}, \mathbf{b})$  with an MLE  $(\alpha_0, \beta_0)$ , i.e.,  $\mathbf{r} = \mathbf{a} \otimes \mathbf{1}_k$  and  $\mathbf{c} = \mathbf{b} \otimes \mathbf{1}_k$ , where  $\otimes$  denotes the Kronecker product. Then as  $k \rightarrow \infty$ ,

$$d_{TV}(X_{11}, \mu_{\alpha_0(1) + \beta_0(1)}) = \begin{cases} O\left(k^{-1/2} \sqrt{\log k}\right) & \text{if } \mu = \text{Counting}(\mathbb{Z}_{\geq 0}) \text{ or } \text{Leb}(\mathbb{R}_{\geq 0}) \\ O\left(k^{-1/4} \log k\right) & \text{if } \mu = \text{Wtd versions of the above} \end{cases}$$

- Answers Barvinok's 2010 conjecture on marginal distribution of random contingency tables

## Theorem (Scaling limit in cut norm; L-M '24+)

Let  $\delta$ -tame margins  $(\mathbf{r}_m, \mathbf{c}_n) \rightarrow (\mathbf{r}, \mathbf{c})$  in  $L^1$ . Let  $X \sim \mu^{\otimes(m \times n)}$  cond. on  $X \in \mathcal{T}(\mathbf{r}_m, \mathbf{c}_n)$ . If  $d\mu = h d\text{Counting}(x)$  or  $h dx$  for "nice"  $h$ , with probability at least  $1 - \exp(-C(m\sqrt{n} - n\sqrt{m})(\log(m+n))^2)$ ,

$$\|W_X - W^{\mathbf{r}, \mathbf{c}}\|_{\square} \leq C\sqrt{n^{-1/2} + m^{-1/2}} \log(m+n) + C\sqrt{\|(\mathbf{r}, \mathbf{c}) - (\bar{\mathbf{r}}_m, \bar{\mathbf{c}}_n)\|_1}.$$

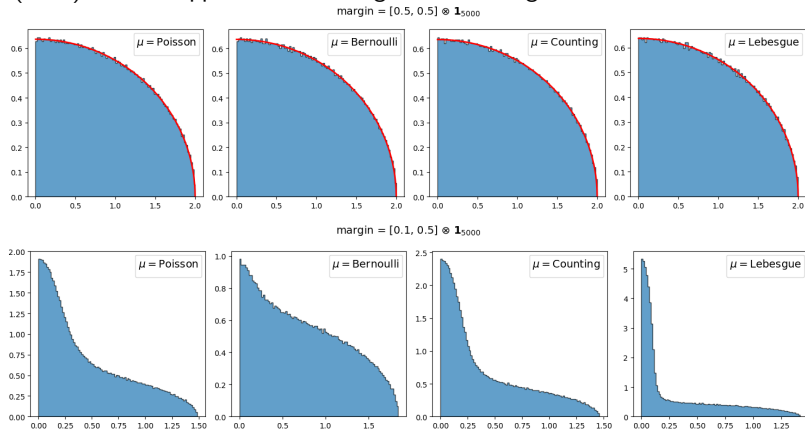
- ▶ We now know  $X \approx \mathbb{E}[Y] = Z^{\mathbf{r}, \mathbf{c}}$ . What about the fluctuation  $X - Z^{\mathbf{r}, \mathbf{c}}$ ?



- ▶ We now know  $X \approx \mathbb{E}[Y] = Z^{\mathbf{r}, \mathbf{c}}$ . What about the fluctuation  $X - Z^{\mathbf{r}, \mathbf{c}}$ ?
- ▶ Since  $X$  is a non-symmetric rectangular RM, we look at the empirical singular value distribution (ESD).

- ▶ We now know  $X \approx \mathbb{E}[Y] = Z^{\mathbf{r}, \mathbf{c}}$ . What about the fluctuation  $X - Z^{\mathbf{r}, \mathbf{c}}$ ?
- ▶ Since  $X$  is a non-symmetric rectangular RM, we look at the empirical singular value distribution (ESD).
  - If  $X$  is unconditioned, then the ESD follows Marchenko-Pastur quarter-circle law (1967). What happens under margin-conditioning?

- ▶ We now know  $X \approx \mathbb{E}[Y] = \mathcal{Z}^{r,c}$ . What about the fluctuation  $X - \mathcal{Z}^{r,c}$ ?
- ▶ Since  $X$  is a non-symmetric rectangular RM, we look at the empirical singular value distribution (ESD).
  - If  $X$  is unconditioned, then the ESD follows Marchenko-Pastur quarter-circle law (1967). What happens under margin-conditioning?



**Figure:** Empirical singular value distribution for  $X \sim \mu^{\otimes(m \times n)}$  given  $X \in \mathcal{T}(r, c)$

## Corollary (Universality of quarter-circle law for uniform margins; L-M 24+)

Assume uniform margins  $\mathbf{r} = \mathbf{c} = a\mathbf{1}_n$  for some  $a$  in the support of  $\mu$ . Let

$$\tilde{X}_n := \frac{1}{\sqrt{2\psi''(\phi(a))n}}(X - a\mathbf{1}\mathbf{1}^\top).$$

Then the empirical singular value distribution of  $\tilde{X}_n$  converges weakly to the Marchenko-Pastur quarter-circle law  $\frac{1}{\pi}\sqrt{4-x^2}dx$  in probability.

## Corollary (Universality of quarter-circle law for uniform margins; L-M 24+)

Assume uniform margins  $\mathbf{r} = \mathbf{c} = a\mathbf{1}_n$  for some  $a$  in the support of  $\mu$ . Let

$$\tilde{X}_n := \frac{1}{\sqrt{2\psi''(\phi(a))n}}(X - a\mathbf{1}\mathbf{1}^\top).$$

Then the empirical singular value distribution of  $\tilde{X}_n$  converges weakly to the Marchenko-Pastur quarter-circle law  $\frac{1}{\pi}\sqrt{4-x^2}dx$  in probability.

- Chatterjee, Diaconis, Sly in 2010 showed the above for  $\mu = \text{Leb}(\mathbb{R}_{\geq 0})$ .

## Corollary (Universality of quarter-circle law for uniform margins; L-M 24+)

Assume uniform margins  $\mathbf{r} = \mathbf{c} = a\mathbf{1}_n$  for some  $a$  in the support of  $\mu$ . Let

$$\tilde{X}_n := \frac{1}{\sqrt{2\psi''(\phi(a))n}}(X - a\mathbf{1}\mathbf{1}^\top).$$

Then the empirical singular value distribution of  $\tilde{X}_n$  converges weakly to the Marchenko-Pastur quarter-circle law  $\frac{1}{\pi}\sqrt{4-x^2}dx$  in probability.

- Chatterjee, Diaconis, Sly in 2010 showed the above for  $\mu = \text{Leb}(\mathbb{R}_{\geq 0})$ .
- Our result confirms universality of the M-P law for constant margin-conditioned RMs.

## Corollary (Universality of quarter-circle law for uniform margins; L-M 24+)

Assume uniform margins  $\mathbf{r} = \mathbf{c} = a\mathbf{1}_n$  for some  $a$  in the support of  $\mu$ . Let

$$\tilde{X}_n := \frac{1}{\sqrt{2\psi''(\phi(a))n}}(X - a\mathbf{1}\mathbf{1}^\top).$$

Then the empirical singular value distribution of  $\tilde{X}_n$  converges weakly to the Marchenko-Pastur quarter-circle law  $\frac{1}{\pi}\sqrt{4-x^2}dx$  in probability.

- Chatterjee, Diaconis, Sly in 2010 showed the above for  $\mu = \text{Leb}(\mathbb{R}_{\geq 0})$ .
- Our result confirms universality of the M-P law for constant margin-conditioned RMs.
- We have a general result for arbitrary  $\delta$ -tame margins. The limiting law is **not** always M-P; determined by the variance profile  $\psi''(\alpha \oplus \beta)$  through QVE determining the Stieltjes transform.

## Corollary (Universality of quarter-circle law for uniform margins; L-M 24+)

Assume uniform margins  $\mathbf{r} = \mathbf{c} = a\mathbf{1}_n$  for some  $a$  in the support of  $\mu$ . Let

$$\tilde{X}_n := \frac{1}{\sqrt{2\psi''(\phi(a))n}}(X - a\mathbf{1}\mathbf{1}^\top).$$

Then the empirical singular value distribution of  $\tilde{X}_n$  converges weakly to the Marchenko-Pastur quarter-circle law  $\frac{1}{\pi}\sqrt{4-x^2}dx$  in probability.

- Chatterjee, Diaconis, Sly in 2010 showed the above for  $\mu = \text{Leb}(\mathbb{R}_{\geq 0})$ .
- Our result confirms universality of the M-P law for constant margin-conditioned RMs.
- We have a general result for arbitrary  $\delta$ -tame margins. The limiting law is **not** always M-P; determined by the variance profile  $\psi''(\alpha \oplus \beta)$  through QVE determining the Stieltjes transform.
- Sketch of Proof:  $\text{ESD}(\tilde{X}_n) \approx \text{ESD}(\tilde{Y}_n)$  by transference;  $\tilde{Y}_n \tilde{Y}_n^*$  generalized Wishart with variance profile  $\psi''(\alpha_n \oplus \beta_n)$ .



Introduction

Random graphs with given degree sequences

A parametric approach for RMs with given margin

Contingency tables and Typical tables

A non-parametric approach to RMs with given margin

Some results on RMs with exactly given margins

**Phase diagram of tame margins**

Open problems

Sinkhorn algorithm

Static Shrödinger bridge

- ▶ All our results so far depends on the margin  $(\mathbf{r}, \mathbf{c})$  being  $\delta$ -tame: i.e., MLE  $(\boldsymbol{\alpha}, \boldsymbol{\beta})$  exists and its entries satisfy

$$A_\delta := \max \left( A + \delta, -\frac{1}{\delta} \right) \leq \psi'(\boldsymbol{\alpha} \oplus \boldsymbol{\beta}) \leq \min \left( B - \delta, \frac{1}{\delta} \right) =: B_\delta.$$

- ▶ All our results so far depends on the margin  $(\mathbf{r}, \mathbf{c})$  being  $\delta$ -tame: i.e., MLE  $(\boldsymbol{\alpha}, \boldsymbol{\beta})$  exists and its entries satisfy

$$A_\delta := \max \left( A + \delta, -\frac{1}{\delta} \right) \leq \psi'(\boldsymbol{\alpha} \oplus \boldsymbol{\beta}) \leq \min \left( B - \delta, \frac{1}{\delta} \right) =: B_\delta.$$

- ▶ Tameness is an implicit condition. Seek for explicit conditions only using extreme values in the margin.

- ▶ All our results so far depends on the margin  $(\mathbf{r}, \mathbf{c})$  being  $\delta$ -tame: i.e., MLE  $(\boldsymbol{\alpha}, \boldsymbol{\beta})$  exists and its entries satisfy

$$A_\delta := \max \left( A + \delta, -\frac{1}{\delta} \right) \leq \psi'(\boldsymbol{\alpha} \oplus \boldsymbol{\beta}) \leq \min \left( B - \delta, \frac{1}{\delta} \right) =: B_\delta.$$

- ▶ Tameness is an implicit condition. Seek for explicit conditions only using extreme values in the margin.
- ▶ For each point  $(s, t) \in (A, B)^2$ ,  $s \leq t$ , we ask if an arbitrary  $m \times n$  margin  $(\mathbf{r}, \mathbf{c})$  satisfying

$$s \leq \mathbf{r}(i)/n, \mathbf{c}(j)/m \leq t \quad \text{for all } (i, j) \in [m] \times [n]$$

is  $\delta$ -tame for some  $\delta = \delta(\mu, s, t) > 0$  depending only on  $\mu$ ,  $s$ , and  $t$ .

- ▶ All our results so far depends on the margin  $(\mathbf{r}, \mathbf{c})$  being  $\delta$ -tame: i.e., MLE  $(\boldsymbol{\alpha}, \boldsymbol{\beta})$  exists and its entries satisfy

$$A_\delta := \max \left( A + \delta, -\frac{1}{\delta} \right) \leq \psi'(\boldsymbol{\alpha} \oplus \boldsymbol{\beta}) \leq \min \left( B - \delta, \frac{1}{\delta} \right) =: B_\delta.$$

- ▶ Tameness is an implicit condition. Seek for explicit conditions only using extreme values in the margin.
- ▶ For each point  $(s, t) \in (A, B)^2$ ,  $s \leq t$ , we ask if an arbitrary  $m \times n$  margin  $(\mathbf{r}, \mathbf{c})$  satisfying

$$s \leq \mathbf{r}(i)/n, \mathbf{c}(j)/m \leq t \quad \text{for all } (i, j) \in [m] \times [n]$$

is  $\delta$ -tame for some  $\delta = \delta(\mu, s, t) > 0$  depending only on  $\mu$ ,  $s$ , and  $t$ .

- ▶ Let  $\Omega(\mu) \subseteq (A, B)^2$  denote the set of all such points  $(s, t)$ .

- ▶ All our results so far depends on the margin  $(\mathbf{r}, \mathbf{c})$  being  $\delta$ -tame: i.e., MLE  $(\boldsymbol{\alpha}, \boldsymbol{\beta})$  exists and its entries satisfy

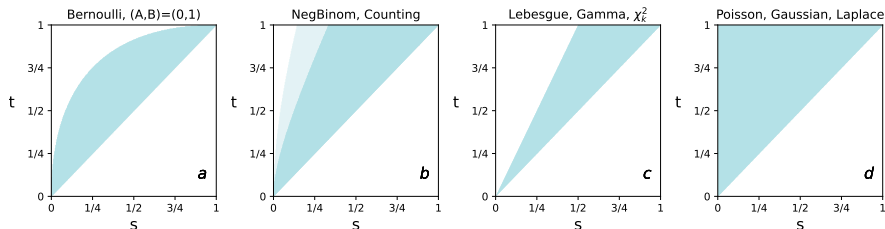
$$A_\delta := \max \left( A + \delta, -\frac{1}{\delta} \right) \leq \psi'(\boldsymbol{\alpha} \oplus \boldsymbol{\beta}) \leq \min \left( B - \delta, \frac{1}{\delta} \right) =: B_\delta.$$

- ▶ Tameness is an implicit condition. Seek for explicit conditions only using extreme values in the margin.
- ▶ For each point  $(s, t) \in (A, B)^2$ ,  $s \leq t$ , we ask if an arbitrary  $m \times n$  margin  $(\mathbf{r}, \mathbf{c})$  satisfying

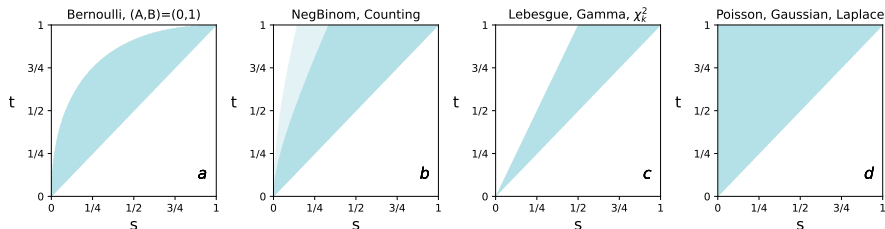
$$s \leq \mathbf{r}(i)/n, \mathbf{c}(j)/m \leq t \quad \text{for all } (i, j) \in [m] \times [n]$$

is  $\delta$ -tame for some  $\delta = \delta(\mu, s, t) > 0$  depending only on  $\mu$ ,  $s$ , and  $t$ .

- ▶ Let  $\Omega(\mu) \subseteq (A, B)^2$  denote the set of all such points  $(s, t)$ .
- ▶ Can we obtain the full phase diagram  $\Omega(\mu)$  for each base measure  $\mu$ ?



**Figure:** Phase diagrams for tame margins for various base measures  $\mu$ . The upper contours are given by  $(s+t)^2 < 2s$ ,  $t \leq 1 + \sqrt{1 + rs^{-1}}$  ( $r = 5$  for NegBinom and  $r = 1$  for Counting),  $t \leq s/2$ , and  $t = \infty$  from left to right.



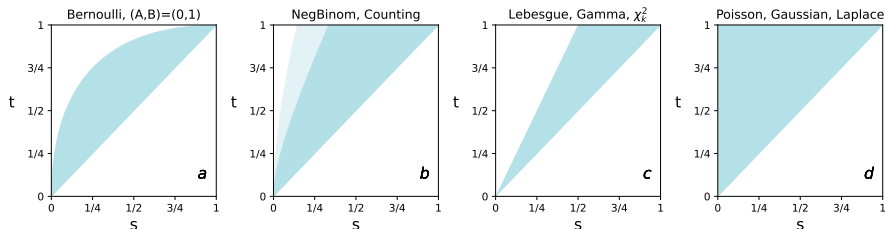
**Figure:** Phase diagrams for tame margins for various base measures  $\mu$ . The upper contours are given by  $(s+t)^2 < 2s$ ,  $t \leq 1 + \sqrt{1 + rs^{-1}}$  ( $r = 5$  for NegBinom and  $r = 1$  for Counting),  $t \leq s/2$ , and  $t = \infty$  from left to right.

## Theorem (L-M '24+)

Suppose  $-\infty < A \leq B < \infty$ . Then each  $(s, t) \in (A, B)^2$  with  $s \leq t$  belongs to  $\Omega(\mu)$  if

$$(s + t - 2A)^2 < 4(B - A)(s - A).$$





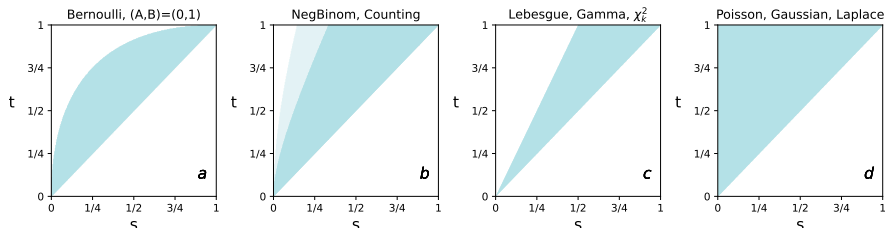
**Figure:** Phase diagrams for tame margins for various base measures  $\mu$ . The upper contours are given by  $(s+t)^2 < 2s$ ,  $t \leq 1 + \sqrt{1 + rs^{-1}}$  ( $r = 5$  for NegBinom and  $r = 1$  for Counting),  $t \leq s/2$ , and  $t = \infty$  from left to right.

## Theorem (L-M '24+)

Suppose  $-\infty < A \leq B < \infty$ . Then each  $(s, t) \in (A, B)^2$  with  $s \leq t$  belongs to  $\Omega(\mu)$  if

$$(s + t - 2A)^2 < 4(B - A)(s - A).$$

- A Matrix version of Erdős-Galai condition



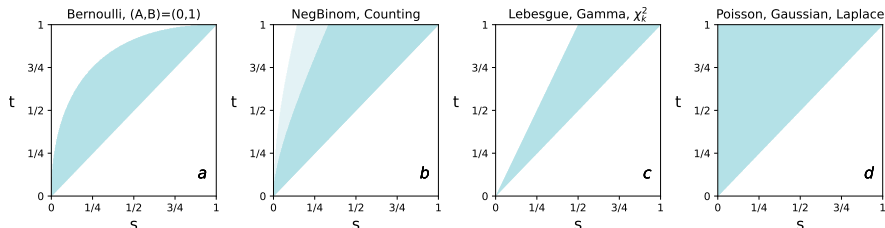
**Figure:** Phase diagrams for tame margins for various base measures  $\mu$ . The upper contours are given by  $(s+t)^2 < 2s$ ,  $t \leq 1 + \sqrt{1 + rs^{-1}}$  ( $r = 5$  for NegBinom and  $r = 1$  for Counting),  $t \leq s/2$ , and  $t = \infty$  from left to right.

## Theorem (L-M '24+)

Suppose  $-\infty < A \leq B < \infty$ . Then each  $(s, t) \in (A, B)^2$  with  $s \leq t$  belongs to  $\Omega(\mu)$  if

$$(s + t - 2A)^2 < 4(B - A)(s - A).$$

- A Matrix version of Erdős-Galai condition
- Universality: only depends on  $B - A$



**Figure:** Phase diagrams for tame margins for various base measures  $\mu$ . The upper contours are given by  $(s + t)^2 < 2s$ ,  $t \leq 1 + \sqrt{1 + rs^{-1}}$  ( $r = 5$  for NegBinom and  $r = 1$  for Counting),  $t \leq s/2$ , and  $t = \infty$  from left to right.

## Theorem (L-M '24+)

$(s, t) \in \Omega(\delta)$  if and only if  $t/s < \lambda_c$  where

$$\lambda_c := \begin{cases} 1 + \sqrt{1 + rs^{-1}} & \text{if } \mu = r\text{-fold convolution of the counting measure on } \mathbb{Z}_{\geq 0}, \\ 2 & \text{if } \mu = \text{Gamma, Lebesgue}(\mathbb{R}_{\geq 0}), \text{ or } \chi_k^2 \\ \infty & \text{if } \mu = \text{Poisson, Gaussian, Laplace} \end{cases}$$

Introduction

Random graphs with given degree sequences

A parametric approach for RMs with given margin

Contingency tables and Typical tables

A non-parametric approach to RMs with given margin

Some results on RMs with exactly given margins

Phase diagram of tame margins

**Open problems**

Sinkhorn algorithm

Static Shrödinger bridge

► (Summary)

If we condition  $X \sim \mu^{\otimes(m \times n)}$  on being in  $\mathcal{T}(r, c)$ , how does it look like?

- $X_{11} \stackrel{d}{=} \mu_{\alpha(1)+\beta(1)}$
- $\mathbb{E}[X] \approx Z^{r,c} = \psi'(\alpha \oplus \beta) = \mathbb{E}[Y]$
- $X - \mathbb{E}[Y] =$  small w.h.p. in cut norm
- $X \stackrel{d}{\approx} Y \sim \mu_{\alpha \oplus \beta}$

► (Summary)

If we condition  $X \sim \mu^{\otimes(m \times n)}$  on being in  $\mathcal{T}(r, c)$ , how does it look like?

- $X_{11} \stackrel{d}{=} \mu_{\alpha(1)+\beta(1)}$
  - $\mathbb{E}[X] \approx Z^{r,c} = \psi'(\alpha \oplus \beta) = \mathbb{E}[Y]$
  - $X - \mathbb{E}[Y] =$  small w.h.p. in cut norm
  - $X \stackrel{d}{\approx} Y \sim \mu_{\alpha \oplus \beta}$
- Empirical distribution of eigenvalues of  $X - \mathbb{E}[Y]$ ? (circular for constant margin with  $\mu = \text{Leb}(\mathbb{R}_{\geq 0})$  by Nguyen '14 [16])

► (Summary)

If we condition  $X \sim \mu^{\otimes(m \times n)}$  on being in  $\mathcal{T}(r, c)$ , how does it look like?

- $X_{11} \stackrel{d}{=} \mu_{\alpha(1)+\beta(1)}$
  - $\mathbb{E}[X] \approx Z^{r,c} = \psi'(\alpha \oplus \beta) = \mathbb{E}[Y]$
  - $X - \mathbb{E}[Y] =$  small w.h.p. in cut norm
  - $X \stackrel{d}{\approx} Y \sim \mu_{\alpha \oplus \beta}$
- Empirical distribution of eigenvalues of  $X - \mathbb{E}[Y]$ ? (circular for constant margin with  $\mu = \text{Leb}(\mathbb{R}_{\geq 0})$  by Nguyen '14 [16])
- Eigenvalue distribution with additional symmetry  $X^\top = X$ ? (Semi-circle for constant margin, generalized Wigner matrices, quadratic vector eq. etc. ; Discussing with Hongchang Ji)

► (Summary)

If we condition  $X \sim \mu^{\otimes(m \times n)}$  on being in  $\mathcal{T}(r, c)$ , how does it look like?

- $X_{11} \stackrel{d}{=} \mu_{\alpha(1)+\beta(1)}$
  - $\mathbb{E}[X] \approx Z^c = \psi'(\alpha \oplus \beta) = \mathbb{E}[Y]$
  - $X - \mathbb{E}[Y] =$  small w.h.p. in cut norm
  - $X \stackrel{d}{\approx} Y \sim \mu_{\alpha \oplus \beta}$
- Empirical distribution of eigenvalues of  $X - \mathbb{E}[Y]$ ? (circular for constant margin with  $\mu = \text{Leb}(\mathbb{R}_{\geq 0})$  by Nguyen '14 [16])
- Eigenvalue distribution with additional symmetry  $X^\top = X$ ? (Semi-circle for constant margin, generalized Wigner matrices, quadratic vector eq. etc. ; Discussing with Hongchang Ji)
- Extend the theory for more general base measure for RM ensemble than i.i.d.? (with  $\mu = \text{Poisson}(1)$ , connection to entropic optimal transport; Ongoing work with William Powell)



► (Summary)

If we condition  $X \sim \mu^{\otimes(m \times n)}$  on being in  $\mathcal{T}(r, c)$ , how does it look like?

- $X_{11} \stackrel{d}{=} \mu_{\alpha(1)+\beta(1)}$
  - $\mathbb{E}[X] \approx Z^{r,c} = \psi'(\alpha \oplus \beta) = \mathbb{E}[Y]$
  - $X - \mathbb{E}[Y] =$  small w.h.p. in cut norm
  - $X \stackrel{d}{\approx} Y \sim \mu_{\alpha \oplus \beta}$
- Empirical distribution of eigenvalues of  $X - \mathbb{E}[Y]$ ? (circular for constant margin with  $\mu = \text{Leb}(\mathbb{R}_{\geq 0})$  by Nguyen '14 [16])
- Eigenvalue distribution with additional symmetry  $X^\top = X$ ? (Semi-circle for constant margin, generalized Wigner matrices, quadratic vector eq. etc. ; Discussing with Hongchang Ji)
- Extend the theory for more general base measure for RM ensemble than i.i.d.? (with  $\mu = \text{Poisson}(1)$ , connection to entropic optimal transport; Ongoing work with William Powell)
- DLP? (For random graphs with given degree sequence, LDP is done by Dhara and Sen '22 [6])
- Ongoing work with Sumit Mukherjee

Thank you very much!

- [1] Alexander Barvinok. “Asymptotic estimates for the number of contingency tables, integer flows, and volumes of transportation polytopes”. In: *International Mathematics Research Notices* 2009.2 (2009), pp. 348–385.
- [2] Alexander Barvinok. “What does a random contingency table look like?” In: *Combinatorics, Probability and Computing* 19.4 (2010), pp. 517–539.
- [3] Petter Brändén, Jonathan Leake, and Igor Pak. “Lower bounds for contingency tables via Lorentzian polynomials”. In: *Israel Journal of Mathematics* 253.1 (2023), pp. 43–90.
- [4] E Rodney Canfield and Brendan D McKay. “Asymptotic enumeration of integer matrices with large equal row and column sums”. In: *Combinatorica* 30.6 (2010), p. 655.
- [5] Sourav Chatterjee, Persi Diaconis, and Allan Sly. “Random graphs with a given degree sequence”. In: *The Annals of Applied Probability* 21.4 (2011), pp. 1400–1435.
- [6] Souvik Dhara and Subhabrata Sen. “Large deviation for uniform graphs with given degrees”. In: *Ann. Appl. Probab.* 32.3 (2022), pp. 2327–53.

- [7] Persi Diaconis and Anil Gangolli. “Rectangular arrays with fixed margins”. In: *Discrete probability and algorithms*. Springer, 1995, pp. 15–41.
- [8] Samuel Dittmer, Hanbaek Lyu, and Igor Pak. “Phase transition in random contingency tables with non-uniform margins”. In: *Transactions of the American Mathematical Society* 373.12 (2020), pp. 8313–8338.
- [9] Robert Fortet. “Résolution d’un système d’équations de M. Schrödinger”. In: *Journal de Mathématiques Pures et Appliquées* 19.1-4 (1940), pp. 83–105.
- [10] IJ Good and JF Crook. “The enumeration of arrays and a generalization related to contingency tables”. In: *Discrete Mathematics* 19.1 (1977), pp. 23–45.
- [11] Irving J Good. “Maximum entropy for hypothesis formulation, especially for multidimensional contingency tables”. In: *The Annals of Mathematical Statistics* 34.3 (1963), pp. 911–934.
- [12] Irving J Good. “On the application of symmetric Dirichlet distributions and their mixtures to contingency tables”. In: *The Annals of Statistics* 4.6 (1976), pp. 1159–1189.
- [13] Isidore Jacob Good. *Probability and the Weighing of Evidence*. C. Griffin London, 1950.

- [14] Catherine Greenhill and Brendan D McKay. “Asymptotic enumeration of sparse nonnegative integer matrices with specified row and column sums”. In: *Advances in Applied Mathematics* 41.4 (2008), pp. 459–481.
- [15] Hanbaek Lyu and Igor Pak. “On the number of contingency tables and the independence heuristic”. In: *Bulletin of the London Mathematical Society* 54.1 (2022), pp. 242–255.
- [16] Hoi H Nguyen. “Random doubly stochastic matrices: the circular law”. In: (2014).
- [17] Michele Pavon, Giulio Trigila, and Esteban G Tabak. “The Data-Driven Schrödinger Bridge”. In: *Communications on Pure and Applied Mathematics* 74.7 (2021), pp. 1545–1573.
- [18] Cédric Villani. *Topics in optimal transportation*. Vol. 58. American Mathematical Soc., 2021.

Introduction

Random graphs with given degree sequences

A parametric approach for RMs with given margin

Contingency tables and Typical tables

A non-parametric approach to RMs with given margin

Some results on RMs with exactly given margins

Phase diagram of tame margins

Open problems

Sinkhorn algorithm

Static Shrödinger bridge

- ▶ Solve the dual (MLE) problem:

- ▶ Solve the dual (MLE) problem:

- $$\arg \max_{\alpha, \beta} \left( g^{\mathbf{r}, \mathbf{c}}(\alpha, \beta) := \langle \mathbf{r}, \alpha \rangle + \langle \mathbf{c}, \beta \rangle - \sum_{i,j} \psi(\alpha(i) + \beta(j)) \right)$$



- Solve the dual (MLE) problem:

- $$\arg \max_{\alpha, \beta} \left( g^{r, c}(\alpha, \beta) := \langle \mathbf{r}, \alpha \rangle + \langle \mathbf{c}, \beta \rangle - \sum_{i, j} \psi(\alpha(i) + \beta(j)) \right)$$
- Strictly concave maximization in two variables  $\alpha, \beta$   
→ **Alternating Maximization!** (a.k.a. Nonlinear Gauss-Seidel or BCD)

$$\begin{cases} \alpha_k \leftarrow \arg \max_{\alpha \in \mathbb{R}^m} g^{r, c}(\alpha, \beta_{k-1}) \\ \beta_k \leftarrow \arg \max_{\beta \in \mathbb{R}^n} g^{r, c}(\alpha_k, \beta). \end{cases}$$

- ▶ Solve the dual (MLE) problem:

- $$\arg \max_{\alpha, \beta} \left( g^{\mathbf{r}, \mathbf{c}}(\alpha, \beta) := \langle \mathbf{r}, \alpha \rangle + \langle \mathbf{c}, \beta \rangle - \sum_{i,j} \psi(\alpha(i) + \beta(j)) \right)$$
- Strictly concave maximization in two variables  $\alpha, \beta$   
→ **Alternating Maximization!** (a.k.a. Nonlinear Gauss-Seidel or BCD)

$$\begin{cases} \alpha_k \leftarrow \arg \max_{\alpha \in \mathbb{R}^m} g^{\mathbf{r}, \mathbf{c}}(\alpha, \beta_{k-1}) \\ \beta_k \leftarrow \arg \max_{\beta \in \mathbb{R}^n} g^{\mathbf{r}, \mathbf{c}}(\alpha_k, \beta). \end{cases}$$

- Finding critical points for the marginal problems, it reduces to

$$\begin{cases} \text{For } 1 \leq i \leq m, \alpha_k(i) \leftarrow \text{unique } \alpha \text{ s.t. } \mathbf{r}(i) = \sum_{j=1}^n \psi'(\alpha + \beta_{k-1}(j)), \\ \text{For } 1 \leq j \leq n, \beta_k(j) \leftarrow \text{unique } \beta \text{ s.t. } \mathbf{c}(j) = \sum_{i=1}^m \psi'(\alpha_k(i) + \beta). \end{cases}$$

- ▶ Solve the dual (MLE) problem:

- $$\arg \max_{\alpha, \beta} \left( g^{\mathbf{r}, \mathbf{c}}(\alpha, \beta) := \langle \mathbf{r}, \alpha \rangle + \langle \mathbf{c}, \beta \rangle - \sum_{i,j} \psi(\alpha(i) + \beta(j)) \right)$$
- Strictly concave maximization in two variables  $\alpha, \beta$   
→ **Alternating Maximization!** (a.k.a. Nonlinear Gauss-Seidel or BCD)

$$\begin{cases} \alpha_k \leftarrow \arg \max_{\alpha \in \mathbb{R}^m} g^{\mathbf{r}, \mathbf{c}}(\alpha, \beta_{k-1}) \\ \beta_k \leftarrow \arg \max_{\beta \in \mathbb{R}^n} g^{\mathbf{r}, \mathbf{c}}(\alpha_k, \beta). \end{cases}$$

- Finding critical points for the marginal problems, it reduces to

$$\begin{cases} \text{For } 1 \leq i \leq m, \alpha_k(i) \leftarrow \text{unique } \alpha \text{ s.t. } \mathbf{r}(i) = \sum_{j=1}^n \psi'(\alpha + \beta_{k-1}(j)), \\ \text{For } 1 \leq j \leq n, \beta_k(j) \leftarrow \text{unique } \beta \text{ s.t. } \mathbf{c}(j) = \sum_{i=1}^m \psi'(\alpha_k(i) + \beta). \end{cases}$$

- For  $\mu = \text{Poisson}(1)$  (Schrödinger bridge),  $\psi'(x) = e^x$ , so

$$\begin{cases} \text{For } 1 \leq i \leq m, \alpha_k(i) \leftarrow \log(\mathbf{r}(i)) - \log\left(\sum_{j=1}^n \exp(\beta_{k-1}(j))\right), \\ \text{For } 1 \leq j \leq n, \beta_k(j) \leftarrow \log(\mathbf{c}(j)) - \log\left(\sum_{i=1}^m \exp(\alpha_k(i))\right). \end{cases}$$

- Solve the dual (MLE) problem:

- $$\arg \max_{\alpha, \beta} \left( g^{\mathbf{r}, \mathbf{c}}(\alpha, \beta) := \langle \mathbf{r}, \alpha \rangle + \langle \mathbf{c}, \beta \rangle - \sum_{i,j} \psi(\alpha(i) + \beta(j)) \right)$$
- Strictly concave maximization in two variables  $\alpha, \beta$   
→ **Alternating Maximization!** (a.k.a. Nonlinear Gauss-Seidel or BCD)

$$\begin{cases} \alpha_k \leftarrow \arg \max_{\alpha \in \mathbb{R}^m} g^{\mathbf{r}, \mathbf{c}}(\alpha, \beta_{k-1}) \\ \beta_k \leftarrow \arg \max_{\beta \in \mathbb{R}^n} g^{\mathbf{r}, \mathbf{c}}(\alpha_k, \beta). \end{cases}$$

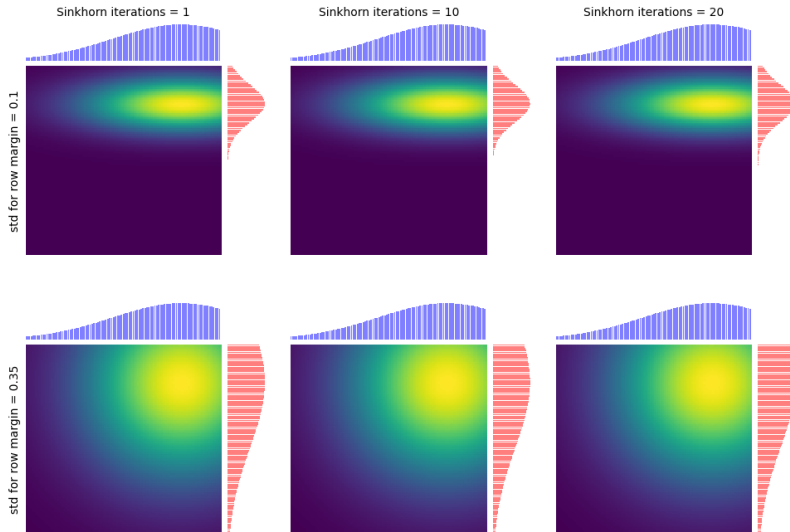
- Finding critical points for the marginal problems, it reduces to

$$\begin{cases} \text{For } 1 \leq i \leq m, \alpha_k(i) \leftarrow \text{unique } \alpha \text{ s.t. } \mathbf{r}(i) = \sum_{j=1}^n \psi'(\alpha + \beta_{k-1}(j)), \\ \text{For } 1 \leq j \leq n, \beta_k(j) \leftarrow \text{unique } \beta \text{ s.t. } \mathbf{c}(j) = \sum_{i=1}^m \psi'(\alpha_k(i) + \beta). \end{cases}$$

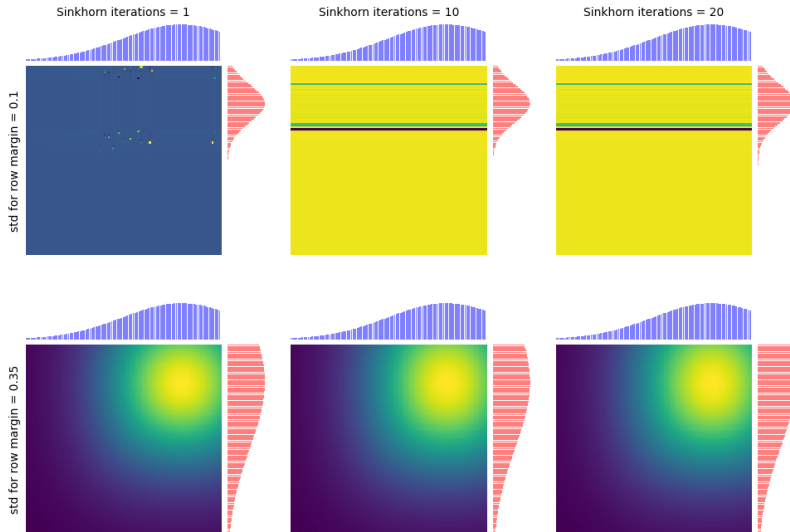
- For  $\mu = \text{Poisson}(1)$  (Schrödinger bridge),  $\psi'(x) = e^x$ , so

$$\begin{cases} \text{For } 1 \leq i \leq m, \alpha_k(i) \leftarrow \log(\mathbf{r}(i)) - \log\left(\sum_{j=1}^n \exp(\beta_{k-1}(j))\right), \\ \text{For } 1 \leq j \leq n, \beta_k(j) \leftarrow \log(\mathbf{c}(j)) - \log\left(\sum_{i=1}^m \exp(\alpha_k(i))\right). \end{cases} \quad \text{Sinkhorn!!}$$

## Poisson typical tables



## Counting typical tables



## Theorem (Linear convergence of generalized Sinkhorn; L-Mukherjee '24+ )

Fix  $\mu$  arbitrary. Let  $(\alpha_k, \beta_k)$  = generalised Sinkhorn iterates. Fix an MLE  $(\alpha^*, \beta^*)$  for  $\delta$ -tame  $(\mathbf{r}, \mathbf{c})$  and denote  $\Delta_k := g^{\mathbf{r}, \mathbf{c}}(\alpha^*, \beta^*) - g^{\mathbf{r}, \mathbf{c}}(\alpha_k, \beta_k)$ . Suppose  $\psi''$  is monotonic and  $\alpha_0 = \mathbf{0}$  or  $\mu$  admits arbitrary tilting ( $\Theta = \mathbb{R}$ ). Then

$$\frac{\sigma_-(\varepsilon)^2}{2} \|(\alpha^* \oplus \beta^*) - (\alpha_k \oplus \beta_k)\|_F^2 \leq \Delta_k \leq \left(1 - \frac{\sigma_-(\varepsilon)^4}{\sigma_+(\varepsilon)^4}\right)^{k-1} \Delta_1 \quad \text{for all } k \geq 1.$$

## Theorem (Linear convergence of generalized Sinkhorn; L-Mukherjee '24+ )

Fix  $\mu$  arbitrary. Let  $(\alpha_k, \beta_k)$  = generalised Sinkhorn iterates. Fix an MLE  $(\alpha^*, \beta^*)$  for  $\delta$ -tame  $(\mathbf{r}, \mathbf{c})$  and denote  $\Delta_k := g^{\mathbf{r}, \mathbf{c}}(\alpha^*, \beta^*) - g^{\mathbf{r}, \mathbf{c}}(\alpha_k, \beta_k)$ . Suppose  $\psi''$  is monotonic and  $\alpha_0 = \mathbf{0}$  or  $\mu$  admits arbitrary tilting ( $\Theta = \mathbb{R}$ ). Then

$$\frac{\sigma_-(\varepsilon)^2}{2} \|(\alpha^* \oplus \beta^*) - (\alpha_k \oplus \beta_k)\|_F^2 \leq \Delta_k \leq \left(1 - \frac{\sigma_-(\varepsilon)^4}{\sigma_+(\varepsilon)^4}\right)^{k-1} \Delta_1 \quad \text{for all } k \geq 1.$$

► Key Challenges:



## Theorem (Linear convergence of generalized Sinkhorn; L-Mukherjee '24+ )

Fix  $\mu$  arbitrary. Let  $(\alpha_k, \beta_k) =$  generalised Sinkhorn iterates. Fix an MLE  $(\alpha^*, \beta^*)$  for  $\delta$ -tame  $(\mathbf{r}, \mathbf{c})$  and denote  $\Delta_k := g^{\mathbf{r}, \mathbf{c}}(\alpha^*, \beta^*) - g^{\mathbf{r}, \mathbf{c}}(\alpha_k, \beta_k)$ . Suppose  $\psi''$  is monotonic and  $\alpha_0 = \mathbf{0}$  or  $\mu$  admits arbitrary tilting ( $\Theta = \mathbb{R}$ ). Then

$$\frac{\sigma_-(\varepsilon)^2}{2} \|(\alpha^* \oplus \beta^*) - (\alpha_k \oplus \beta_k)\|_F^2 \leq \Delta_k \leq \left(1 - \frac{\sigma_-(\varepsilon)^4}{\sigma_+(\varepsilon)^4}\right)^{k-1} \Delta_1 \quad \text{for all } k \geq 1.$$

### ► Key Challenges:

- **The set of MLEs is unbounded:**  $(\alpha^*, \beta^*)$  MLE  $\iff (\alpha^* + \lambda, \beta^* - \lambda)$  MLE  $\forall \lambda \in \mathbb{R}$

## Theorem (Linear convergence of generalized Sinkhorn; L-Mukherjee '24+ )

Fix  $\mu$  arbitrary. Let  $(\alpha_k, \beta_k)$  = generalised Sinkhorn iterates. Fix an MLE  $(\alpha^*, \beta^*)$  for  $\delta$ -tame  $(\mathbf{r}, \mathbf{c})$  and denote  $\Delta_k := g^{\mathbf{r}, \mathbf{c}}(\alpha^*, \beta^*) - g^{\mathbf{r}, \mathbf{c}}(\alpha_k, \beta_k)$ . Suppose  $\psi''$  is monotonic and  $\alpha_0 = \mathbf{0}$  or  $\mu$  admits arbitrary tilting ( $\Theta = \mathbb{R}$ ). Then

$$\frac{\sigma_-(\varepsilon)^2}{2} \|(\alpha^* \oplus \beta^*) - (\alpha_k \oplus \beta_k)\|_F^2 \leq \Delta_k \leq \left(1 - \frac{\sigma_-(\varepsilon)^4}{\sigma_+(\varepsilon)^4}\right)^{k-1} \Delta_1 \quad \text{for all } k \geq 1.$$

### ► Key Challenges:

- The set of MLEs is unbounded:  $(\alpha^*, \beta^*)$  MLE  $\iff (\alpha^* + \lambda, \beta^* - \lambda)$  MLE  $\forall \lambda \in \mathbb{R}$
- Need a priori bound on the Sinkhorn iterates

## Theorem (Linear convergence of generalized Sinkhorn; L-Mukherjee '24+ )

Fix  $\mu$  arbitrary. Let  $(\alpha_k, \beta_k)$  = generalised Sinkhorn iterates. Fix an MLE  $(\alpha^*, \beta^*)$  for  $\delta$ -tame  $(\mathbf{r}, \mathbf{c})$  and denote  $\Delta_k := g^{\mathbf{r}, \mathbf{c}}(\alpha^*, \beta^*) - g^{\mathbf{r}, \mathbf{c}}(\alpha_k, \beta_k)$ . Suppose  $\psi''$  is monotonic and  $\alpha_0 = \mathbf{0}$  or  $\mu$  admits arbitrary tilting ( $\Theta = \mathbb{R}$ ). Then

$$\frac{\sigma_-(\varepsilon)^2}{2} \|(\alpha^* \oplus \beta^*) - (\alpha_k \oplus \beta_k)\|_F^2 \leq \Delta_k \leq \left(1 - \frac{\sigma_-(\varepsilon)^4}{\sigma_+(\varepsilon)^4}\right)^{k-1} \Delta_1 \quad \text{for all } k \geq 1.$$

### ► Key Challenges:

- The set of MLEs is unbounded:  $(\alpha^*, \beta^*)$  MLE  $\iff (\alpha^* + \lambda, \beta^* - \lambda)$  MLE  $\forall \lambda \in \mathbb{R}$
- Need a priori bound on the Sinkhorn iterates  
 $\Leftarrow$  For Schrödinger bridge ( $\mu = \text{Poisson}(1)$ ), use exact form of Sinkhorn updates

## Theorem (Linear convergence of generalized Sinkhorn; L-Mukherjee '24+ )

Fix  $\mu$  arbitrary. Let  $(\alpha_k, \beta_k)$  = generalised Sinkhorn iterates. Fix an MLE  $(\alpha^*, \beta^*)$  for  $\delta$ -tame  $(\mathbf{r}, \mathbf{c})$  and denote  $\Delta_k := g^{\mathbf{r}, \mathbf{c}}(\alpha^*, \beta^*) - g^{\mathbf{r}, \mathbf{c}}(\alpha_k, \beta_k)$ . Suppose  $\psi''$  is monotonic and  $\alpha_0 = \mathbf{0}$  or  $\mu$  admits arbitrary tilting ( $\Theta = \mathbb{R}$ ). Then

$$\frac{\sigma_-(\varepsilon)^2}{2} \|(\alpha^* \oplus \beta^*) - (\alpha_k \oplus \beta_k)\|_F^2 \leq \Delta_k \leq \left(1 - \frac{\sigma_-(\varepsilon)^4}{\sigma_+(\varepsilon)^4}\right)^{k-1} \Delta_1 \quad \text{for all } k \geq 1.$$

### ► Key Challenges:

- The set of MLEs is unbounded:  $(\alpha^*, \beta^*)$  MLE  $\iff (\alpha^* + \lambda, \beta^* - \lambda)$  MLE  $\forall \lambda \in \mathbb{R}$
- Need a priori bound on the Sinkhorn iterates  
 $\Leftarrow$  For Schrödinger bridge ( $\mu = \text{Poisson}(1)$ ), use exact form of Sinkhorn updates  
 For general  $\mu$ , no exact form of Sinkhorn updates (implicit)

## Theorem (Linear convergence of generalized Sinkhorn; L-Mukherjee '24+ )

Fix  $\mu$  arbitrary. Let  $(\alpha_k, \beta_k) =$  generalised Sinkhorn iterates. Fix an MLE  $(\alpha^*, \beta^*)$  for  $\delta$ -tame  $(\mathbf{r}, \mathbf{c})$  and denote  $\Delta_k := g^{\mathbf{r}, \mathbf{c}}(\alpha^*, \beta^*) - g^{\mathbf{r}, \mathbf{c}}(\alpha_k, \beta_k)$ . Suppose  $\psi''$  is monotonic and  $\alpha_0 = \mathbf{0}$  or  $\mu$  admits arbitrary tilting ( $\Theta = \mathbb{R}$ ). Then

$$\frac{\sigma_-(\varepsilon)^2}{2} \|(\alpha^* \oplus \beta^*) - (\alpha_k \oplus \beta_k)\|_F^2 \leq \Delta_k \leq \left(1 - \frac{\sigma_-(\varepsilon)^4}{\sigma_+(\varepsilon)^4}\right)^{k-1} \Delta_1 \quad \text{for all } k \geq 1.$$

### ► Key Challenges:

- The set of MLEs is unbounded:  $(\alpha^*, \beta^*)$  MLE  $\iff (\alpha^* + \lambda, \beta^* - \lambda)$  MLE  $\forall \lambda \in \mathbb{R}$

- Need a priori bound on the Sinkhorn iterates

$\Leftarrow$  For Schrödinger bridge ( $\mu = \text{Poisson}(1)$ ), use exact form of Sinkhorn updates  
For general  $\mu$ , no exact form of Sinkhorn updates (implicit)

- Solution: We show the  $\ell^\infty$ -distance between the iterates and the set of MLEs does not expand

Introduction

Random graphs with given degree sequences

A parametric approach for RMs with given margin

Contingency tables and Typical tables

A non-parametric approach to RMs with given margin

Some results on RMs with exactly given margins

Phase diagram of tame margins

Open problems

Sinkhorn algorithm

Static Shrödinger bridge

- ▶ Given a base probability measure  $\mathcal{R}$  on  $\mathbb{R}^2$  and marginal distributions  $\mu_1$  and  $\mu_2$ ,

$$(**) \quad \min_{\mathcal{H} \in \Pi(\mu_1, \mu_2)} D_{KL}(\mathcal{H} \parallel \mathcal{R})$$

The optimal  $\mathcal{H}$  from above is the **static Schrödinger bridge** between  $\mu_1$  and  $\mu_2$  w.r.t.  $\mathcal{R}$  [9, 17]

- ▶ Given a base probability measure  $\mathcal{R}$  on  $\mathbb{R}^2$  and marginal distributions  $\mu_1$  and  $\mu_2$ ,

$$(**) \quad \min_{\mathcal{H} \in \Pi(\mu_1, \mu_2)} D_{KL}(\mathcal{H} \parallel \mathcal{R})$$

The optimal  $\mathcal{H}$  from above is the **static Schrödinger bridge** between  $\mu_1$  and  $\mu_2$  w.r.t.  $\mathcal{R}$  [9, 17]

- ▶  $\exists \alpha_1, \alpha_2 : \mathbb{R} \rightarrow \mathbb{R}$ , the **Schrödinger potentials** [18] s.t.

$$\frac{d\mathcal{H}}{d\mathcal{R}}(x, y) = e^{\alpha_1(x) + \alpha_2(y)} \quad \mathcal{R}\text{-a.s.}$$



- ▶ Given a base probability measure  $\mathcal{R}$  on  $\mathbb{R}^2$  and marginal distributions  $\mu_1$  and  $\mu_2$ ,

$$(**) \quad \min_{\mathcal{H} \in \Pi(\mu_1, \mu_2)} D_{KL}(\mathcal{H} \parallel \mathcal{R})$$

The optimal  $\mathcal{H}$  from above is the **static Schrödinger bridge** between  $\mu_1$  and  $\mu_2$  w.r.t.  $\mathcal{R}$  [9, 17]

- ▶  $\exists \alpha_1, \alpha_2 : \mathbb{R} \rightarrow \mathbb{R}$ , the **Schrödinger potentials** [18] s.t.

$$\frac{d\mathcal{H}}{d\mathcal{R}}(x, y) = e^{\alpha_1(x) + \alpha_2(y)} \quad \mathcal{R}\text{-a.s.}$$

- ▶ Specializing  $\mathcal{R} \propto e^{-\gamma/\varepsilon} \mu_1 \otimes \mu_2$ ,  $(**)$  becomes **Entropic Optimal Transport**:

$$\min_{\mathcal{H} \in \Pi(\mu_1, \mu_2)} \int_{\mathbb{R}^2} \gamma(x, y) \mathcal{H}(dx, dy) + \varepsilon D_{KL}(\mathcal{H} \parallel \mu_1 \otimes \mu_2),$$

- ▶ Given a base probability measure  $\mathcal{R}$  on  $\mathbb{R}^2$  and marginal distributions  $\mu_1$  and  $\mu_2$ ,

$$(**) \quad \min_{\mathcal{H} \in \Pi(\mu_1, \mu_2)} D_{KL}(\mathcal{H} \parallel \mathcal{R})$$

The optimal  $\mathcal{H}$  from above is the **static Schrödinger bridge** between  $\mu_1$  and  $\mu_2$  w.r.t.  $\mathcal{R}$  [9, 17]

- ▶  $\exists \alpha_1, \alpha_2 : \mathbb{R} \rightarrow \mathbb{R}$ , the **Schrödinger potentials** [18] s.t.

$$\frac{d\mathcal{H}}{d\mathcal{R}}(x, y) = e^{\alpha_1(x) + \alpha_2(y)} \quad \mathcal{R}\text{-a.s.}$$

- ▶ Specializing  $\mathcal{R} \propto e^{-\gamma/\varepsilon} \mu_1 \otimes \mu_2$ , **(\*\*)** becomes **Entropic Optimal Transport**:

$$\min_{\mathcal{H} \in \Pi(\mu_1, \mu_2)} \int_{\mathbb{R}^2} \gamma(x, y) \mathcal{H}(dx, dy) + \varepsilon D_{KL}(\mathcal{H} \parallel \mu_1 \otimes \mu_2),$$

- ▶ Specializing  $\mathcal{R} = \text{Uniform}([m] \times [n])$ , **(\*\*)** becomes

$$\min_{X=(x_{ij}) \in (0, \infty)^{m \times n}} x_{ij} \log x_{ij} \quad \text{subject to} \quad \sum_{j=1}^n x_{ij} = \mu_1(i), \quad \sum_{i=1}^m x_{ij} = \mu_2(j) \quad \forall i, j.$$

- ▶ Given a base probability measure  $\mathcal{R}$  on  $\mathbb{R}^2$  and marginal distributions  $\mu_1$  and  $\mu_2$ ,

$$(**) \quad \min_{\mathcal{H} \in \Pi(\mu_1, \mu_2)} D_{KL}(\mathcal{H} \parallel \mathcal{R})$$

The optimal  $\mathcal{H}$  from above is the **static Schrödinger bridge** between  $\mu_1$  and  $\mu_2$  w.r.t.  $\mathcal{R}$  [9, 17]

- ▶  $\exists \alpha_1, \alpha_2 : \mathbb{R} \rightarrow \mathbb{R}$ , the **Schrödinger potentials** [18] s.t.

$$\frac{d\mathcal{H}}{d\mathcal{R}}(x, y) = e^{\alpha_1(x) + \alpha_2(y)} \quad \mathcal{R}\text{-a.s.}$$

- ▶ Specializing  $\mathcal{R} \propto e^{-\gamma/\varepsilon} \mu_1 \otimes \mu_2$ ,  $(**)$  becomes **Entropic Optimal Transport**:

$$\min_{\mathcal{H} \in \Pi(\mu_1, \mu_2)} \int_{\mathbb{R}^2} \gamma(x, y) \mathcal{H}(dx, dy) + \varepsilon D_{KL}(\mathcal{H} \parallel \mu_1 \otimes \mu_2),$$

- ▶ Specializing  $\mathcal{R} = \text{Uniform}([m] \times [n])$ ,  $(**)$  becomes

$$\min_{X=(x_{ij}) \in (0, \infty)^{m \times n}} x_{ij} \log x_{ij} \quad \text{subject to} \quad \sum_{j=1}^n x_{ij} = \mu_1(i), \quad \sum_{i=1}^m x_{ij} = \mu_2(j) \quad \forall i, j.$$

This is in fact the **typical table** problem with  $\mu = \text{Poisson}(1)$ !

- $x_{ij} \log x_{ij} = D(\mu_{\phi(x_{ij})} \parallel \mu) + x_{ij} - 1$